

ENDOGENOUS INSTITUTIONS AND ECONOMIC OUTCOMES.*

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Abstract

In this paper, I evaluate the relative importance of a “culture of cooperation,” understood as the implicit reward from cooperating in prisoner’s dilemma and investment types of activities, and “inclusive political institutions,” which enable the citizenry to check the executive authority. To elaborate, I divide Europe into 120 km \times 120 km grid cells, and I exploit exogenous variation in both institutions driven by persistent medieval history. In particular, I document strong first-stage relationships between present-day norms of respect and trust and the severity of consumption risk—i.e., climate volatility—over the 1000-1600 period and between present-day regional political autonomy and the factors that raised the returns on elite-citizenry investments, i.e., the terrain ruggedness and the direct access to the coast. Using this instrumental variables approach, I show that only culture has a first order effect on development, even after controlling for country fixed effects, medieval innovations, the present-day role of medieval geography, and the factors modulating the impact of institutions. Crucially, the excluded instruments have no direct impact on development, and the effect of culture holds within pairs of adjacent grid cells with different medieval climate volatility. An explanation for these results is that culture, but not a more inclusive political process, is necessary to produce public-spirited politicians and push voters to punish political malfeasance. Micro-evidence from Italian Parliament data supports this idea.

Keywords: Geography; Culture; Democracy; Development; Political Accountability.

JEL classification: Z10; H10; O10; D72.

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1 Introduction

Huge empirical evidence suggests that a “culture of cooperation,” understood as the implicit reward from cooperating in prisoner’s dilemma and investment types of activities, and “inclusive political institutions,” which enable the citizenry to check the executive authority, foster economic development and are correlated with past inclusive political institutions (Tabellini, 2010; Guiso et al., 2016). Documenting however that the two types of arrangements reinforce one another and are persistent does not help identify their relative importance. This paper tackles this issue by devising a multiple instrumental variables approach that exploits exogenous variation in both present-day culture of cooperation and inclusive political institutions created at the European regional level by persistent medieval history.

From the 11th century on indeed, the lords started to offer the peasants high powered farming contracts to exploit the improved land productivity and to enter into commercial partnerships with a rising class of merchants engaged in the first long-distance trades. These innovations flourished where the lords gave up some political power to gain credibility as investment partners and persisted where the citizenry also credibly committed to cooperate in investment by attracting the Cistercians and Franciscans. Both monastic orders dictated a culture of cooperation in exchange for guidance on how to share consumption risk and spread where the climate was very—but not too—erratic. Inspired by these facts and previous related research (Fleck and Hanssen, 2006; Durante, 2010),¹ Boranbay and Guerriero (2016) employ a panel of 90 European regions spanning the 1000-1600 period to test the following two ideas. First, the prospect of a sufficiently profitable investment pushes the elite to introduce more inclusive political institutions to convince the citizens that a sufficient part of its return will be shared via public spending. Second, the citizenry accumulates culture to share consumption risk and credibly commit to cooperate while investing with the elite. The “commitment dimension of cultural accumulation” also reduces the elite’s temptation to repeal political reforms after a fall of the investment value. Consistent with these predictions, medieval reforms toward tighter constraints on the elite’s power are positively driven by the

¹While Fleck and Hanssen (2006) show that in Ancient Greece democratization was stronger where the elite found more difficult to monitor the citizens’ farming investments, Durante (2010) documents that Europeans living today in regions in which the climate was more erratic between 1500 and 1750 trust more others.

factors determining the observability and thus the profitability of farming investments and by the value of long-distance trades, i.e., respectively the ruggedness of the terrain and the direct access to the coast. In addition, cultural accumulation, as captured by the discounted number of years Cistercian and Franciscan houses were active per square km, rises with the risk of harvest destruction, as driven by the volatility of the growing season temperature, and with shocks depressing the investment value, i.e., the opening of the Atlantic routes. Since present-day institutions are deeply rooted into medieval ones, the correlation between past institutional arrangements, created by the commitment dimension of cultural accumulation, produces first-stage relationships between past political infrastructures and both present-day culture and political institutions. These however are not distinct and therefore insufficient to identify the relative present-day importance of the two institutions.

To deal with these issues, I devise a multiple instrumental variables approach exploiting the geographic determinants of past institutions. The success of this identification strategy depends on the power of the two sets of instruments isolating the role of each institution. Operationally, I divide Europe into $120 \text{ km} \times 120 \text{ km}$ grid cells, and I show that the volatility of the 1000-1600 growing season temperature has a strong effect on present-day culture of cooperation, as captured by the strength of norms of respect and trust self-reported to the 2008 European Value Study, and no impact on a measure of the inclusiveness of regional political institutions averaged between 1950 and 2010. This is obtained supplementing the constraints on the executive authority score developed by the Polity IV project with information on the political autonomy from the central government of the NUTS 2 regions in the sample. The latter has been recognized by a large literature as a key determinant of the citizenry's ability to monitor politicians (Frey, 2005), and it displays a strong within-country correlation with regional measures of property rights protection proposed by Charron et al. (2014). The terrain ruggedness and the direct access to the coast instead have a large impact on current political institutions and a little effect on present-day culture. Building on these separate first-stages, I show that only culture has a first order effect on the natural logarithm of the GDP per capita averaged between 2002 and 2009, even after controlling for country fixed effects, medieval innovations, intermediate outcomes, factors modulating the impact of permanent institutions, and the present-day role of medieval geography, i.e., present-day cli-

mate volatility, average distance from the sea, and average within-grid traveling distance. In particular, these last three observables control for the possible stickiness of climate volatility and the effect of the excluded instruments on present-day tourism and trade.

Including this rich conditioning set makes difficult to envision that the excluded instruments might have a direct impact on current outcomes via channels other than permanent institutions and, in particular, through their present-day economic role. To address the concern of whether indeed the exclusion restriction is satisfied, I perform a number of robustness tests. First, I document that the overidentifying restrictions cannot be rejected at a level nowhere lower than 75% conditional on all observables and that the excluded instruments have no direct impact on outcomes in the semi-reduced form regressions. Second, I show that the effect of culture on economic success survives within pairs of neighboring grid cells differing in their medieval climate volatility. This exercise enables me to control for all unobservables specific to the $120 \text{ km} \times 240 \text{ km}$ grid cell-pairs. Finally, I perform the following falsification test to examine the reduced form relationship between the volatility of the 1000-1600 growing season temperature and present-day GDP per capita inside and outside my sample. Within Europe, I find a strong positive link between the two variables as expected, given my two-stage least squares—2SLS from here on—estimates. Regions, which experienced a more erratic weather and thus accumulated a stronger culture by attracting more Cistercians and Franciscans, are more developed today. If medieval climate volatility affects income only through a persistent risk-sharing-driven culture of cooperation, I should not find a similar relationship where the cost of accumulating past culture was prohibitive because of the opposition to Western monasticism. This is what I find. Building on 117 Turkish grid cells, I estimate a statistically insignificant relationship between medieval climate volatility and present-day income. This is consistent with the barriers to Western monasticism erected there by the Eastern Orthodox Church first and the Ottoman empire then.

An explanation for these results is that more inclusive political institutions are irrelevant in facilitating the monitoring of politicians by voters if the latter are not morally compelled to punish political malfeasance or if the former have weak civic virtues (Boix and Posner, 1998; Padró i Miquel et al., 2015). To confirm this idea, I show that there are fewer criminal prosecutions of Italian Parliament members in electoral districts in which culture is stronger

but not in those endowed with more inclusive political institutions (see also Nannicini et al., [2013]). This evidence points at a key mechanism inducing the primacy of culture.

The papers most closely related to mine are Acemoglu and Johnson (2005) and Tabellini (2010). The latter also tries to overcome problems inherent to cross-country data by focusing on a cross-section of 69 European regions and using past political institutions as excluded instrument for present-day culture in growth regressions. Differently from this and the related contributions on the within-country effect of past institutions (Michalopoulos and Papaioannou, 2013; Gennaioli et al., 2013; Di Liberto and Sideri, 2015), I devise a strategy dealing explicitly with the possibility that within-country confounding factors may drive at the same time past institutions, present-day institutions, and present-day outcomes. Acemoglu and Johnson (2005) instead share with me the aim of unbundling institutions but focus on contract enforcement and property rights. No previous study however has identified the separate roles of culture and inclusive political institutions. Crucially, I do so by exploiting their geographic determinants in a sample in which geography has neither shaped present-day economies through persistent innovations nor modulated the spread of slavery (Nunn and Puga, 2012) and the colonizers' settlement strategy (Acemoglu et al., 2001).

The paper proceeds as follows. I illustrate the key historical facts about the medieval institutional revolution in section 2. Next, I describe the data and the empirical strategy in section 3. Then, I assess the relative importance of culture and inclusive political institutions in section 4, and I present the micro-evidence on political accountability in section 5. Finally, I conclude in section 6, and I gather tables and figures in the appendix.

2 The Medieval Origins of European Institutions

The anarchy created by the fall of the Western Roman empire pushed the population to seek the protection of the lords who, empowered by the feudal contract, pacified their estates [Stearns 2001, p. 165-176]. This new order fueled a revolution that changed Europe to date.

Attracted by the prospect of improved land productivity and the opportunity of long-distance trades, the lords began to enter into high-powered farming contracts with the peasants and commercial partnerships with a rising class of merchants, who obtained exemption from the tolls necessary to cross the land [Stearns 2001, p. 191-222]. These contractual

innovations flourished where the lords also introduced more inclusive political institutions to fortify their credibility as investment partners [Stearns 2001, p. 216], and in particular in the *Giudicati* in Sardinia (952-1297), the communes of Northern Italy and France (1080-1282), the maritime republics of Genoa, Pisa, and Venice (1099-1406), the towns of Aragon and Cataluña (1150-1213), the German imperial cities (1152-1806), and the Swiss Cantons (1291-1515). To illustrate, Frederick I granted the communal privileges to the difficult-to-reach Northern Italian communities in exchange for the sizable payments fixed by the 1183 Peace of Constance [Stearns 2001, p. 208], whereas the *communes jurées* of Northern France and the Flanders were chartered by the early Capetian kings interested in gaining from the lucrative exchanges of woolens for Eastern spices [Stearns 2001, p. 199]. Organized as a sworn association of free men and governed by a public assembly selecting the executive, these states were “aimed at economic prosperity [and favored by the lord’s] immediate political and financial considerations” [Stearns 2001, p. 199].

Meanwhile, Western monasticism was transforming interpersonal relationships. Imported from the East during the 5th century, it spread out across Europe through some ascetic and lots of lax initiatives until a group of dissatisfied Cluniac monks abandoned Molesme in Burgundy and founded in 1098 a new monastery in Cîteaux [Burton and Kerr 2011, p. 9-10]. This event opened a new and highly influential phase of the medieval Church. The Cistercians indeed revived the original Benedictine emphasis on poverty, prayer, and manual labor to diffuse the novel and powerful idea, illustrated in their 1119 *Carta Caritatis*, that both the partnership between monasteries and the interaction between worshipers should be rooted in “mutual love and esteem, combined with a benevolent eye to human frailty [i.e.,] charity rather than the exercise of power” [Tobin 1995, p. 40]. Crucially, these charity-based norms of conduct should not materialize through alms but through cooperation [Burton and Kerr 2011, p. 28-29], which the Cistercians themselves supported by organizing a series of risk-sharing activities with the help of local laypeople known as *conversi* and secular laborers [Burton and Kerr 2011, p. 150-163; Donkin 1978, p. 39]. First, they accepted as grants mainly undeveloped lands and turned them into fertile compact holdings disseminating at the same time advanced farming techniques [Donkin 1978, p. 172-173; Tobin 1995, p. 43]. Initially targeted at rendering the neighboring villages self-sufficient, with the demise of the

conversi system these estates were progressively leased to the peasants at rates lower than those set by the lords [Donkin 1978, p. 111; Burton and Kerr 2011, p. 166]. Second, they further insulated the population from shocks by setting up trade fairs, developing international trade agreements, and diversifying economic activities with the introduction of forges and mills [Tobin 1995, p. 128; Burton and Kerr 2011, p. 185]. Finally, they provided a series of other risk-sharing services, like shelter for those in need and food for the starved, significantly limiting in this way social and religious conflicts [Burton and Kerr 2011, p. 47-50 and 191-194]. These activities, so crucial in a world of risk-minimization, eased the diffusion of the charity-based norms of cooperation the Cistercians championed in the communities first exposed to their action and desperate to preserve it [Burton and Kerr 2011, p. 120]. Moreover, they urged the populations of the neighboring areas—especially those located where the climate was very unpredictable, but not too erratic to force re-siting—to either offer the White monks a site for building a new house or push local houses to join the order [Knowles 1948, p. 64; Donkin 1978, p. 36; Berman 2000, p. 95, 107, and 223; Burton and Kerr 2011, p. 23-36]. The relationship of “kinship” among houses, which was enforced by the duties of cross-visitation and support, assured the homogeneity of the order’s action [Tobin 1995, p. 41; Burton and Kerr 2011, p. 82] and connected regions divided by national conflicts, making “generalized” the Cistercian morality [Burton and Kerr 2011, p. 94]. Not surprisingly, in 1153 there were already 435 Cistercian houses scattered around Europe.

When the 14th century “emancipation of the villein class [...] combined with the visitations of pestilence” [Knowles 1948, p. 77] undermined the *conversi* system, the Cistercians slowly left the scene to the Franciscans [Tobin 1995, p. 125 and 236]. Exactly as the former had “opened the monastic vocation to the agrarian peasantry” [Lawrence 2001, p. 178], the latter embraced the apostolic life of “poverty [,] active preaching mission [...] and example” [Lawrence 2001, p. 247 and 259] prompted by St. Francis in his 1223 *Regula* to offer a rising “town-dwelling laity [...] the idea of the devout life for the laity” [Lawrence 2001, p. 240 and 259], i.e., a life of “charity pursued through moral consideration and practical engagement” [Muzzarelli 2001, p. 115]. Similarly to the Cistercians and uniquely within the remainder of Western monasticism,² the Friars Minor accepted “unenviable sites” [Knowles 1948, p.

²Albeit in 1215 Pope Innocent III imposed on all monastic orders the Cistercian hierarchical structure,

192] to build with the help of the lay brothers part of the “Third Orders” a dense network of houses, linked by a Cistercian-like kinship, and supervise several key risk-sharing—e.g., micro-credit and public health—activities [Muzzarelli 2001, p. 40]. Among these practices, the most noteworthy was to run the *Monte di Pietà*, which accommodated the customers with loans in return for a pledge auctioned if the loan plus an interest payment evaluated at a rate lower than that charged by private bankers—i.e., 3% versus 30%—was not paid [Muzzarelli 2001, p. 205-206]. Summoned by the representatives of those towns more prone to economic shocks and internal unrests [Muzzarelli 2001, p. 11 and 60], the Franciscan preachers would first gather donations [Muzzarelli 2001, p. 24, 60, and 227], then draft the *Monte*’s constitution having in mind “the customers’ material and moral destinies” [Muzzarelli 2001, p. 219], and finally help run the pawnshop [Muzzarelli 2001, p. 243]. In doing so, they subjected the loan issuance to an evaluation of the “morality and social behaviors of the customers” [Muzzarelli 2001, p. 216], i.e., those citizens who required credit to overcome a moment of need [Muzzarelli 2001, p. 166, 170, and 244] and, if helped, would have actively contributed to make “cohabitation more cooperative and fair” [Muzzarelli 2001, p. 41]. Crucially, the *Monte*’s obligation to back up the citizenry-nobility partnerships in the case of liquidity shocks also strengthened the relationship between the two groups [Muzzarelli 2001, p. 193]. Accordingly, the Franciscan penetration in the Mediterranean delayed the return to autocratic regimes after the opening of the Atlantic routes and the consequent fall in the profitability of the Mediterranean trades [Muzzarelli 2001, p. 36 and 228] and democratization of the Reign of England and the Provinces (Acemoglu et al., 2005). Only the Protestant Reformation deprived Western monasticism of its pivotal role by stigmatizing ecclesiastic property and professional preaching [Tobin 1995, p. 158].

In the following, I exploit this rich natural historical experiment to assess the relative importance of present-day culture and inclusive political institutions in Europe.

3 Data and Empirical Strategy

The sample consists of 578 grid cells in 16 European countries for which I have sufficient

Benedictines, Cluniacs, and Dominicans (Augustinians, Carmelites, Carthusians, Cathars, and Waldensians) specialized instead in theological studies and university teaching (contemplation), possibly accepting lay brothers only as a support for the daily organization of the monastery (Knowles, 1948; Lawrence 2001).

information (see footnote 16 and table 1).³ The grid cells have a 1° width, which is the spatial resolution of the excluded instrument for which I observe the most refined data. Contrary to a region-based approach, this design allows me to compare units of similar size, sidestep the endogeneity of regional boundaries, and exploit a substantial within-country variation.⁴

3.1 Measuring a Culture of Cooperation

The proxy for present-day culture is obtained from the 2008 European Value Study, which in turn is the only wave listing the NUTS 2 region where the respondent lived when 14 and thus culturally mature (Tabellini, 2008; Andersen et al., 2016).⁵ In particular, I follow Boranbay and Guerriero (2016), and I rely on the self-reported extent of “generalized” respect and trust for others. Both norms are meant as abstract rules of cooperative conduct applied outside the reference group of friends and relatives, and so a generalized instead of a “limited” form of morality (Platteau, 2000), which thus captures the Cistercian and Franciscan concept of “Caritas” discussed in section 2 and, more generally, the implicit reward from cooperating in any prisoner’s dilemma and investment type of activities.

To elaborate, respectful individuals are more reluctant to free ride on others and more willing to participate in joint partnerships and politics (Tabellini, 2010). Similarly, trust not only favors cooperation in prisoner’s dilemma games as documented by a broad experimental evidence (Durante, 2010), but it also reduces transaction costs, expands market exchange, and facilitates the division of labor (Dixit, 2004). In order to contemporaneously capture both norms, I construct *Culture* as the first principal component extracted from the share of answers mentioning “tolerance and respect for other people” as important qualities children should be encouraged to learn—i.e., *Respect*—and the share of answers “most people can be trusted” to the question “generally speaking, would you say that most people can be trusted or that you can’t be too careful in dealing with people?”—i.e., *Trust*. If a grid cell belongs to multiple NUTS 2 regions, I assign it a figure equal to the average of the values

³Grid cells located on the borders are divided in units each entirely belonging to a single country. Considering the undivided grid cells to deal with unobserved determinants of national boundaries produces similar results.

⁴Using as cross-section identifiers the regions considered by Boranbay and Guerriero (2016) reduces the average within-country standard deviation in the medieval climate volatility (ruggedness) proxy from 0.05 to 0.04 degree Celsius (0.09 to 0.06 km) and makes the estimates very noisy (see the Internet appendix).

⁵NUTS 2 regions are defined by Eurostat on the basis of administrative criteria and have a population ranging from 800,000 to three million. The average (median) number of respondents per region is 313 (167).

the aforementioned first principal component assumes in each represented region weighted by the region relative contribution to the grid cell land area (see table 2 for a summary of each variable). I follow the same procedure for the other variables measured at the regional level. Tabellini (2010) also considers the convictions that effort is likely to pay off—i.e., *Control*—and the refusal of hierarchical control—i.e., a low level of *Obedience*—as norms conducive to development. Neither of the two however is strictly connected to the norms of cooperation disseminated by the Cistercians and the Franciscans. A legacy of cross-cultural psychology has indeed documented that *Control* mainly concerns “the desirability of individuals independently pursuing their own ideas [and *Obedience*] refers to a cultural emphasis on obeying role obligations within a legitimately unequal distribution of power” [Licht et al. 2007, p. 115]. The gist of the analysis will however be the same, should I turn to either the first principal component extracted from *Respect*, *Trust*, *Control*, and *Obedience*—i.e., *Culture-T*—or one between *Respect* and *Trust* (see the Internet appendix).

The upper-right map in figure 1 illustrates the large variation in *Culture* across Europe and the size of the grid cells I use below as cross-section identifiers relative to the NUTS 2 regions surveyed by the European Value Study. Even if continuous measures are used in the empirical analysis, data are displayed in the maps in five intervals whose break points are chosen to best group similar values and maximize the differences between groups.⁶ Darker colors correspond to higher values. While the Benelux, England, France, Northern Italy, and Northern Spain exhibit the strongest culture of cooperation, the Czech Republic, Poland, and Portugal display the most limited one. As clarified by the comparison between this pattern and that in the upper-left map in figure 1, present-day norms of respect and trust are deeply rooted in the medieval risk-sharing-driven culture of cooperation, which I proxy with the discounted number of years Cistercian and Franciscan houses were active per square km averaged over the 1000-1600 period, i.e., *Culture-M*.⁷ Albeit immaterial to our main

⁶The goodness of variance fit method minimizes the average deviation of the interval values from the interval mean, while maximizing the average deviation of the interval values from the means of the other intervals.

⁷For each of the 684 (2931) Cistercian (Franciscan) houses and each half-century between 1000 and 1600, this figure equals the difference between the number of years in which the house had operated and those elapsed from its possible closure per square km if positive and zero otherwise. To obtain the raw data, I eliminate from the lists of monasteries reported in Van Der Meer (1965) and Moorman (1983) those that are not indicated in at least another of the available sources of their location, foundation, and closure, i.e., <http://www.cistercensi.info/>, <http://users.bart.nl/~roestb/franciscan/>, and the bibliography therein.

findings, the discounting emphasizes the importance of the monks’ activity (see also Persson and Tabellini, [2009]). On the contrary, scaling the years of activity by the region area instead of its population is necessary to correctly represent the two orders’ diffusion since a minimum distance between houses was compulsory [Burton and Kerr 2011, p. 44].

As seen in section 2, both monastic orders dictated norms of respect and trust in exchange for guidance on how to share consumption risk, and under the threat of defecting, to the populations more subject to consumption risk and thus more interested in securing their services. Given the substantial homogeneity of the two orders’ activities and their distinctiveness within medieval Western monasticism, *Culture-M* then gauges the input to the technology that transformed the citizenry’s involvement with culture into evolutionary stable norms, and higher values detect a stronger culture of cooperation in the past.⁸ This interpretation is consistent with two fundamental insights of evolutionary psychology and Malthusian growth theories: a social group dictates to its members, via natural selection and cross-punishment, cultural norms maximizing its fitness (Barkow et al., 1992; Clark, 2005), and these values are stronger the larger the culturally-driven reproductive advantage is (Andersen et al., 2016). Focusing on the Cistercians, Andersen et al. (2016) propose a similar mechanism but describe them as aimed at spreading values of hard work and thrift. Albeit consistent with Baumol (1990) and Weber (1958) himself, this vision is at odds with the more recent and substantial historic literature introduced in section 2. Contrary to what speculated by Andersen et al. (2016) indeed, the fundamental issue distancing the order’s founders from Molesme was not its “failure to observe the Rule of St Benedict [but the fact that it] was rich [and] “association of possession with virtues is not usually long-lasting”⁹ [Burton and Kerr 2011, p. 11]. This reasoning led the Cistercians to embrace a cult of corporate poverty and austerity exemplified in very taxing rules of monastic life and a deep contempt for those members of the community seeking social competition and accumulation of wealth [Burton and Kerr 2011, p. 103-118 and 155-156]. To elaborate, even when man-

⁸To further cross-validate this variable, Boranbay and Guerriero (2016) report its high correlation—0.8—with the number of years the *Monti* were active per square km. Since these pawnshops survived only when loans were repaid [Muzzarelli 2001, p. 189-244], their endurance is positively related to the likelihood of successful risk-sharing activities, and therefore it is an outcome-based measure of past culture just as the electoral turnout and blood donations are of present-day culture (see Guiso et al., [2016]).

⁹As emphatically reported in the *Exordium Parvum*, the Cistercians’ narrative of the order’s origin.

aging market-oriented enterprises, the White monks considered effort and profit as merely instrumental to fund their risk-sharing activities and so fulfill their moralization mission [Burton and Kerr 2011, p. 187]. To begin with, it was the desire of rationalizing neighboring economies, injecting liquidity in unstable markets, and making the lords’ property available to the peasantry to guide the expansion of the order’s holdings [Burton and Kerr 2011, p. 160-168]. Accordingly, it should not strike as strange that several houses experienced an endemic lack of savings, which possibly plunged them into bankruptcy first and either royal custody or abandon later on [Burton and Kerr 2011, p. 174]. Second, the fees and tolls obtained from the organization of fairs were often invested in charitable activities [Donkin 1978, p. 159]. Finally, even when the demise of the *conversi* system made leasing the only market-oriented endeavor viable for the order, these agreements were usually conditioned on the peasants’ obligation to provide risk-sharing services [Burton and Kerr 2011, p. 177]. All in all, while it is very hard to see in the order’s action a desire to support “cultural values [assisting] the rise of capitalism outside the monastic walls” [Andersen et al. 2016, p. 2], it seems natural to interpret it, as the Cistercians did in the *Carta Caritatis*, as their duty “to be of service to [their brothers,] avoid the evil of avarice [and] retain the care of their souls for the sake of charity.” Similar conclusions can be drawn for the Franciscans whose life indeed “demanded not only exterior imitation of Christ through poverty [...] but also interior conformity through self-denial, obedience, humility, and love” [Daniel 1992, p. 46].

Consistent with the above discussion, I show that none among the share of answers to the 2008 European Value Study mentioning “hard work” as an important quality that children should be encouraged to learn—*Hard-Work*, the share of answers reporting “thrift”—*Thrift*, *Control*, and *Obedience* is positively and significantly correlated with either *Culture-M* or medieval climate volatility (see the Internet appendix). Moreover, when the proxy for culture is the first principal component extracted from *Hard-Work*, *Thrift*, *Respect*, and *Trust*—i.e., *Culture-A*, the message of the empirical analysis is pretty similar (see the Internet appendix).

3.2 Measuring the Inclusiveness of Political Institutions

I define the inclusiveness of present-day political institutions as the strength of the rules enabling voters to select more public-spirited representatives and check more closely their

decisions. To capture both aspects,¹⁰ I consider the average over the 1950-2010 period of the sum of the Polity IV constraints on the executive authority score and a regional political autonomy index, i.e., *Democracy*. The political autonomy index takes value 1 if the region had exclusive control over a limited set of policy—e.g., education, 2 if it was also fiscally decentralized, 3 if it had substantial political autonomy from the central government,¹¹ and 0 otherwise. Conditional on fixed effects, *Democracy* gauges two important sources of institutional variation, i.e., the differences between the autocracy and the democracy that ran respectively Eastern and Western Germany before their unification and the diverging experiences of the autonomous regions of Austria, Belgium, France, Italy, Spain, and the UK. Despite previous contributions have exclusively studied the former aspect (Persson and Tabellini, 2009), a growing body of regional studies highlights the importance of considering the sizable sub-national variation in political autonomy (Charron et al., 2014). Politicians elected in autonomous regions are directly accountable for local policies, are chosen for their fit with the preferences of the local population, and can design public goods fulfilling the most these preferences (Kappeler et al., 2013). Accordingly, the regional political autonomy index I develop displays a strong within-country correlation with regional measures of property rights protection (see the Internet appendix).¹² This dimension, which the extant literature recognizes as crucial in distinguishing differently inclusive political regimes (Acemoglu and Johnson, 2005), will be lost if one relies only on cross-country variation (see also Di Liberto and Sideri, [2015]). My results do not merely rest on the way in which *Democracy* is defined since I obtain similar estimates when I consider only the regional political autonomy index, I focus on the 2000-2010 period, or I use the first principal component extracted from the Polity IV score and the regional political autonomy index (see the Internet appendix).

The bottom-right map in figure 1 displays the considerable variation in *Democracy*. On the one hand, the experience of differently inclusive political regimes has created an institutional gap between the regions located on the two sides of the Iron Curtain. On the other

¹⁰The Polity IV constraints on the executive authority score ranges between one and seven, and higher values indicate stronger constraints on the decision-making power of chief executives.

¹¹I consider a region as fiscally decentralized if it can raise part of its fiscal revenues through region-specific taxes and spend them on local public goods. I treat a unit as politically autonomous if it is fiscally decentralized, can elect its own parliament, and controls all policies except those of national relevance like defense.

¹²These are a measure of honesty, impartiality, and quality of law enforcement, one of the overall quality of governance, and an inverse metrics of the relevance of corruption (Charron et al., 2014).

hand, South Tyrol, Région Wallonne, Vlaams Gewest, Corse, the Italian and Spanish regions, Northern Ireland, Scotland, and Wales have been entrusted in the postwar period a more or less complete political autonomy by their central governments. These arrangements range from the exclusive legislative power on specific matters, like education, granted to all Italian regions in 2001 (Article 117, Italian Constitution) to the almost complete autonomy obtained by the linguistic areas of Belgium in 1962 and the devolved UK regions in 1999. In these last cases, the central governments have kept their responsibility for excepted matters like defense, whereas the regional Parliaments have acquired the residual legislative power and the ability to invest regional tax revenues into local public goods. All in all, *Democracy* ranges from a minimum of 3.20 scored by the ex-Eastern Germany regions of Brandenburg and Sachsen to a maximum of 9 observed, for instance, in Vlaams-Brabant.

The bottom-left map in figure 1 prompts that the present-day heterogeneity in regional political institutions has its roots in medieval history. This map depicts the average over the 11th-16th centuries of the constraints on the elite’s power score coded by Boranbay and Guerriero (2016) for each half-century between 1000 and 1600, *Democracy-M*. The score is obtained by first merging those neighboring NUTS 2 administrative units that, according to Sellier and Sellier (2002), were part of the same state for most of the period (see table 1) and then analyzing the history of each of the medieval states in a 40-year window around each date. During the Middle Ages, the most intense democratization processes were experienced by the agrarian communities of Aragon and Cataluña, the commercial “Giudicati” of Sardinia and communes of Northern Italy, and the maritime republics first and by the Provinces and the Reign of England then. In the postwar period, the autonomy of these states has been partially restored with the justification that the preferences for public good of a historically homogeneous community should be satisfied by local representatives (Frey, 2005).

3.3 Empirical Strategy

The options open to a society characterized by a weak culture of cooperation but more inclusive political institutions are very different from those left to a society in which the political process is less democratic but cooperation is facilitated by solid norms of respect and trust. While the former can barely sustain decentralized markets, investment, and

division of labor (Putnam et al., 1993; Dixit, 2004; Boranbay and Guerriero, 2016), the latter has always the option of relying on informal networks enforcing contracts and protecting property rights (Greif, 2006). Moreover, culture shapes the way citizens participate in policy-making and the behaviors of public officials. On the one hand, it reduces the citizens’ cost of punishing political malfeasance by relaxing collective action constraints, building their qualities of judgment, and shifting their preferences toward community-oriented policies (Boix and Posner, 1998; Padró i Miquel et al., 2015). On the other hand, inconsiderate public officials are likely to engage in nepotism and corruption even in the face of “de jure democratic institutions” (Putnam et al., 1993). The very unequal achievements of the public administration and the judiciary in Northern and Southern Italy despite the 150 years of common political trajectory constitute a glaring example (de Oliveira and Guerriero, 2016).

Hence, it is reasonable to suppose that the performance of a region characterized by a forceful culture but less inclusive political institutions—e.g., Emilia Romagna—will be superior to that of a region in which a more democratic political process is left in the hands of less respectful citizens, e.g., Sardinia. In the following, I show how this observation helps make sense of the comparison between culture and inclusive political institutions.

3.3.1 Unbundling Institutions

Lacking sufficient exogenous variation to identify nonlinearities,¹³ I focus on the equation

$$Y_{i,c} = \alpha_c + \beta_0 C_{i,c} + \gamma_0 D_{i,c} + \delta'_0 \mathbf{X}_{i,c} + \epsilon_{i,c}, \quad (1)$$

where $Y_{i,c}$ is the natural logarithm of the GDP per capita in grid cell i of country c , in euro, averaged between 2002 and 2009, i.e., *Income*.¹⁴ Its source is Eurostat, which collects the data at the NUTS 2 regional level. I obtain similar results if I switch to the G-Econ estimate of the GDP per capita in 1985, which is available at the 1° spatial resolution (see the Internet appendix). $C_{i,c}$ and $D_{i,c}$ denote *Culture* and *Democracy* respectively, and $\mathbf{X}_{i,c}$ gathers the latitude and longitude of the centroid of the grid cell—i.e., *Latitude* and *Longitude*—and

¹³When the three excluded instruments are used to identify $C_{i,c}$, $D_{i,c}$ and their interaction, they become weak.

¹⁴Given the seesawing performance of some European regions, it would be more instructive to link the medieval institutional revolution to the development of each grid cell over a longer spell of time. Unfortunately, to the best of my knowledge, the only proxies for $Y_{i,c}$ at the regional level are those I consider.

possibly the controls discussed below. α_c takes into account country-wide unobservables relevant for development like the legacy of past wars (Iyigun et al., 2015), legal origins (Guerriero, 2016a), and genetic diversity (Ashraf and Galor, 2013).¹⁵ Since the correlation between $C_{i,c}$ and $D_{i,c}$ is only 0.26, multicollinearity is not an issue.

The simplest strategy is to estimate equation (1) by OLS. There are two key issues with this strategy. First, both *Culture* and *Democracy* are endogenous, so I may capture reverse causality or the effect of an omitted variable like real and human capital or religious beliefs. Second, both variables are measured with error, so there may be a downward attenuation bias. To evaluate these concerns, I compare the inconsistent OLS estimates with those obtained by using 2SLS with distinct excluded instruments for *Culture* and *Democracy*. These should be correlated with the endogenous regressors but orthogonal to any omitted variable, i.e., uncorrelated with the dependent variable through any channel other than the endogenous regressors. This strategy should take care of the reverse causality and omitted variable biases as well as of the differential measurement errors in the two endogenous regressors, as long as the measurement errors have the classical form and β_0 and γ_0 can be consistently estimated (see Acemoglu and Johnson, [2005]). The two first-stage regressions are

$$\begin{aligned} C_{i,c} &= \alpha_c + \zeta_1 T_{i,c} + \eta_1 R_{i,c} + \theta_1 I_{i,c} + \delta'_1 \mathbf{X}_{i,c} + \omega_{i,c}, \\ D_{i,c} &= \alpha_c + \zeta_2 T_{i,c} + \eta_2 R_{i,c} + \theta_2 I_{i,c} + \delta'_2 \mathbf{X}_{i,c} + \nu_{i,c}, \end{aligned} \quad (2)$$

where $T_{i,c}$ is the volatility of the 1000-1600 growing season temperature and corresponds to the excluded instrument for culture (see section 3.3.2). $R_{i,c}$ and $I_{i,c}$ label respectively the terrain ruggedness and a dummy for direct access to the coast and represent instead the excluded instruments for inclusive political institutions (see section 3.3.2). The exclusion restriction is that in the population $Cov(\epsilon_{i,c}, T_{i,c}) = Cov(\epsilon_{i,c}, R_{i,c}) = Cov(\epsilon_{i,c}, I_{i,c}) = 0$.

In judging the adequacy of my empirical strategy, two remarks should be heeded. First, using past institutions as excluded instruments does not unbundle present-day institutions because of the commitment dimension of cultural accumulation. As Boranbay and Guerriero (2016) show, there is a strong correlation between the activity of Cistercians and Franciscans and the inclusiveness of political institutions in the Middle Ages since cultural accumulation

¹⁵Failing to account for these confounding factors makes the estimates very noisy (see the Internet appendix).

by the population served as a commitment device when the value of investments fell, and so the elite was tempted to repeal political reforms, i.e., the Franciscans' spread in the Mediterranean intensified after the opening of the Atlantic routes. Therefore, the stickiness of institutions produces first-stages that are not distinct. Moreover, they are also weaker than those detailed in equation (2) since past institutions are measured with error. Accordingly, this different approach delivers estimates that are similar but more noisy than those discussed below (see the Internet appendix). Second, there could be a non zero correlation among $\epsilon_{i,c}$, $\omega_{i,c}$, and $\nu_{i,c}$. Accordingly, I compare the 2SLS results with those obtained estimating equations (1) and (2) as a system by three-stage least squares, 3SLS from here on.

3.3.2 The Geographic Determinants of Medieval Institutions

Building on the historical events illustrated in section 2, Boranbay and Guerriero (2016) study accumulation of culture and democratization in a simple and yet general society. Formally, “elite” members and “citizens” can either share consumption risk with any other individual or invest with a member of a different group. While the first activity resembles a prisoner-dilemma interaction and gauges a more fundamental form of cooperation aimed at hedging against consumption shocks, the second more profitable one captures a more advanced form of cooperation producing a taxable value, e.g., long-distance trades. First, each group costly instills into its members a psychological gain from cooperating, for instance, by attracting a monastic order. This implicit reward embodies a culture of cooperation. Next, the elite selects the political regime. Democracy allows the citizenry to fix the share of investment value to be spent on the production of a public good and its type, whereas autocracy gives these prerogatives to the elite. Then, the agents are randomly matched, and the elite selects the activity if she meets the citizenry. Finally, taxation and public good production follow a cooperative investment. The activity-specific factors—i.e., the severity of consumption risk and the investment value—are exogenous, e.g., geography. Since inefficiencies in public good production render investment infeasible under autocracy, the equilibrium has two key features. While the prospect of a sufficiently profitable investment pushes the elite to enact democracy to convince the citizens that a sufficient part of its return will be shared, accumulation of culture rises with the severity of consumption risk at

its moderate values and then drops at its high values making cheating too appealing.

Consistent with these predictions, culture over the 1000-1600 period and its present-day counterpart are stronger in those European regions in which it was more necessary to cope with consumption risk because of the higher but never extreme climate volatility (Boranbay and Guerriero, 2016). On top of this evidence, I elect as instrument for *Culture* the standard deviation of the 1000-1600 spring-summer temperature in degree Celsius, i.e., *Climate-M*. The raw data are collected from Guiot et al. (2010) and cover most of Europe at the 5° spatial resolution for all the years between 600 and 2000.¹⁶ Each observation is “reconstructed” from indirect proxies such as tree-rings, ice cores, pollens, and indexed climate series based on historical documents. To the best of my knowledge, these are the only gridded data of the pre-1500 European climate. If grid cell i belongs to multiple climatic grids, I assign this grid cell a figure equal to the average of the values medieval climate volatility assumes in each represented grid weighted by the grid relative contribution to grid cell i land area. Allowing clustering by country to account for the within-country correlation in the error term produced by the climate data resolution implies similar second-stages but weakens the first-stages (see the Internet appendix). The same happens when I deal with generic spatial dependence in the error term by turning to the Conley’s (1999) standard errors (see the Internet appendix). Higher resolution gridded data on temperature and rainfall have been devised for the post-1500 period building mainly on instrumental sources (Durante, 2010). Since however pre-1800 climate stations are scarce, these series are much less accurate than reconstructed data and so cannot be incorporated into the analysis (Guiot et al., 2010).¹⁷

The exposition so far suggests that the most relevant feature that could undermine the exclusion restriction is the persistent impact of the economic progress that was triggered, together with the institutional revolution, by medieval geography. To illustrate, since *Climate-*

¹⁶Because of data availability (To have sufficient within-country variation), I exclude from the sample part of Ireland, Portugal, Scotland, and Spain (Andorra, Gibraltar, Luxembourg, Malta, and San Marino). This choice has no relevant impact on the estimates. Moreover, I do not consider Scandinavia and the areas east of Poland and Slovakia and south-east of Hungary and Slovenia for two reasons. First, there are insufficient data on the rest of the medieval states to which they belonged. Second, Western monasticism did not propagate there because of the Orthodox Church’s opposition [Tobin 1995, p. 144].

¹⁷Over the 16th century, the average volatility of the Luterbacher et al.’s (2004) measure of the growing season temperature, which is estimated building on instrumental data, is nine times bigger than that of the Guiot et al.’s (2010) reconstructions, which are instead tailored to preserve a meaningful comparison over time.

M is related to medieval development through agricultural productivity and the adoption of the advances in farming technology spread by the Cistercians, it might directly affect present-day outcomes if this progress had enduring consequences. Even if this occurrence seems unlikely given the intrinsically Malthusian structure of medieval economies (Galor, 2011) and the limited importance of the primary sector in the sample,¹⁸ I show that key measures of medieval farming progress do not confound the effect of permanent institutions. The same can be said of present-day climate volatility and those present-day intermediate outcomes most heavily influenced by *Culture-M*, i.e., financial development, human capital, and Catholic beliefs. As a consequence, it is quite difficult to envision that the climate volatility of more than four centuries ago shapes present-day performance through a channel other than a culture of cooperation conditional on country fixed effects, medieval farming progress, present-day climate volatility, and present-day intermediate outcomes.

Boranbay and Guerriero (2016) also document that between 1000 and 1600 reforms toward tighter constraints on the elite’s power were mostly driven by the factors shaping the value of farming and long-distance trade investments. For what concerns the former, the central driver of the medieval agriculture revolution was the adoption of the heavy plow, which required as many as eight oxen to pull it and forced the peasants to combine their ox teams and split their lands into interspersed strips to ensure that everyone got some land plowed (Slocum, 2005). Thus, the elite’s prospective returns on such a complex investment were higher the more difficult were its monitoring and the plowing itself (see for a similar argument Fleck and Hanssen, [2006]). Building on these remarks, I employ as second excluded instrument the terrain ruggedness in km retrieved from the G-Econ project, i.e., *Ruggedness*. Turning to long-distance trade investments, their value was significantly higher if a direct access to the coast was available being terrestrial movements heavily regulated in the Middle Ages (Brady et al., 1994; Acemoglu et al., 2005). Thus, my third excluded instrument is a dummy for direct access to the Mediterranean and/or the Atlantic Ocean, i.e., *Coast*.

Since *Ruggedness* and *Coast* are related respectively to medieval farming and long-distance and in particular Atlantic trades, they might affect *Income* if the advances they fostered are long-lived. Below, I control not only for medieval farming progress, as aforemen-

¹⁸The share of active population employed in the primary sector between 2002 and 2008 was 6% (see table 2).

tioned, but also for the relevance of Atlantic trades, and I show that these innovations do not confound the effect of permanent institutions. The same can be said of the average distance from the sea and the average within-grid traveling distance, and thus *Ruggedness* and *Coast* are not shaping today economies via their present-day impact on tourism and trade. On top of these observations, it is hard to think that *Ruggedness* and *Coast* drive directly *Income* conditional on country fixed effects, medieval innovations, and their present-day role.

A gaze at figures 1 and 2 reveals not only the sizable variation in both institutions and geography but also that the results obtained by Boranbay and Guerriero (2016) hold in my sample. As table 3 shows, both past institutions and their determinants are powerful drivers of present-day institutions, and the coefficients confirm the theoretical predictions.¹⁹

4 Culture Versus Inclusive Political Institutions

A glance at figures 1 and 2 already hints at the main result of the paper. The regional pattern of present-day per capita output in the leftmost map in figure 2 is similar to that of present-day culture in the upper-right map in figure 1 and that of medieval climate volatility in the central map in figure 2. Northern Italy, Western France, and Northern Spain enjoy a higher development, display stronger norms of respect and trust, and experienced a more erratic medieval climate than the rest of the sample. On the contrary, Czech Republic, Western Poland, and Portugal are marked by very low values of all three variables. The correlations among economic outcomes, culture, and medieval climate volatility are however imperfect. England is one of the most culturally and economically advanced European regions but did not face an unpredictable climate during the Middle Ages, whereas Southern Spain exhibits low values of *Income* and *Climate-M* but quite a strong culture of cooperation. Similarly, the relationships among *Ruggedness*, *Coast*, and *Income* are not clear-cut. Although graphical comparisons are instructive, multivariate analysis is more convincing.

4.1 Main Results

Table 4 reports the OLS, 2SLS, and 3SLS estimates of the different specifications of

¹⁹While a series of recent theoretical papers have clarified that cultural norms inherited from earlier generations deeply shape current culture (see Tabellini, [2008]), an expanding body of empirical contributions has highlighted the persistence of political infrastructures (Acemoglu et al., 2001; Di Liberto and Sideri, 2015).

equation (1). A comparison between columns (1) and (2) suggests that OLS underestimate the impact of culture on per capita income. In fact, switching from OLS to 2SLS increases the coefficient on *Culture* from 0.17 to 0.78. On the contrary, the coefficient on *Democracy* remains quite similar in the two columns. This evidence is consistent with the aforementioned idea that the performance of a region, in which the vacuum opened by less inclusive political institutions is filled by informal networks sustained by a forceful culture, will be superior to that of a region in which a more democratic political process is left in the hands of less respectful citizens. To illustrate, the estimate of β_0 in column (2) is significant at 1% and implies that a one-standard deviation rise in *Culture*—i.e., 0.36—will lead to a 28% rise in present-day GDP per capita and that moving from the lowest level of *Culture*—i.e., - 1.49 in the Balearic Islands—to its mean will increase *Income* by 126%. In contrast, the coefficient on *Democracy* in column (2) is not statistically significant. The two upper(bottom)-scatter plots in figure 3 display graphically the OLS (2SLS) estimates in column (1) (column (2)) highlighting quite clearly that they are not driven by a handful of abnormal observations.²⁰

4.2 Robustness and Sensitivity Checks

The basic estimates imply that while culture has a first-order positive effect on income per capita, the impact of more inclusive political institutions albeit positive is small and statistically insignificant. This is consistent with a series of empirical papers concluding that the average effect of democracy is at most weak (see Glaeser et al., [2004]; Persson and Tabellini, [2009]). Next, I illustrate a number of robustness and sensitivity checks.

4.2.1 Controlling for Observables

To ascertain whether the exclusion restriction holds, I include into $\mathbf{X}_{i,c}$ not only the alternative channels through which the excluded instruments could shape *Income*, but also those determinants of development either driven by or affecting institutions. Including these covariates also helps me assess the effective magnitude of the impact of each institution.

Starting with the other channels through which the excluded instruments could affect present-day outcomes, I consider one variable gaging the possibly persistent impact of me-

²⁰Accordingly, I document in the Internet appendix that the evidence remains essentially the same once I keep out from the analysis the outliers spotted with the Cook’s distance (Cook, 1977)

dieval farming progress, one capturing the possibly long-lasting role of Atlantic trade, one accounting for both, and measures of the present-day economic role of the excluded instruments. The first of these covariates is the share of the active population employed in agriculture and fishing averaged between 2002 and 2008 and collected at the NUTS 2 regional level by the Regio project, i.e., *Primary-Sector*. Including this variable into $\mathbf{X}_{i,c}$ takes into account the possibility that medieval climate volatility has influenced the patterns of sectoral specialization. Measures of the relevance of Atlantic trade can be obtained from Acemoglu et al. (2005). Albeit I focus on the number of Atlantic ports active in the grid cell between 1500 and 1850—i.e., *Atlantic-Trade*, the essence of my results will be the same should I turn to the number of potential Atlantic ports between 1500 and 1850 or either the natural logarithm of voyages per year equivalent or the share of total Atlantic trade from the Atlantic ports in the polity to which the grid cell belonged averaged between 1500 and 1850. The last variable I consider is the natural logarithm of the population per square km averaged over the 1000-1600 period, i.e., *LPD-M*. Demographic data are available from Goldewijk et al. (2011) for the 10,000 BC-2000 period and the whole globe at the 5' spatial resolution and are estimated through time-variant allocation algorithms. Since in the Malthusian epoch urbanization corresponded to development (Galor, 2011), *LPD-M* picks other possible effects of medieval farming progress and long-distance trades. I will obtain similar results should I turn to the natural logarithm of either the urbanization rate or the population averaged between 1000 and 1600 and collected from Goldewijk et al. (2011).

I capture the present-day role of medieval geography with: 1. the normalized first principal component extracted from the standard deviation of the temperature in degree Celsius and that of the precipitation in mm both averaged between 1961 and 1990 and retrieved from the G-Econ project, i.e., *Climate*; 2. the average distance to the coast in the grid cell in km as collected from the G-Econ project, i.e., *Distance-to-Coast*; 3. the average traveling distance between the centroid and the corners of the grid cell in km, i.e., *Traveling-Distance*. While including *Climate* tests whether the stickiness of *Climate-M* is directly affecting present-day agriculture and in turn outcomes, considering *Distance-to-Coast* and *Traveling-Distance* allows me to check whether *Coast* and *Ruggedness* are directly determining *Income* by affecting tourism and trade. I get similar estimates if I also consider the temperature in degree Celsius

and the precipitation in mm both averaged over the 1961-1990 period, the land quality for agriculture and its standard deviation, and the grid cell land area (see Internet appendix). These features could affect income by modulating ethnic diversity (Michalopoulos, 2012) and, in turn, the protection of property rights (Guerriero, 2016b).

Medieval institutions and, in particular, culture might affect present-day economies through three major intermediate outcomes. First, since the micro-credit activities introduced by the Franciscans have shaped regional financial markets (Pascali, 2016), I also consider the 2000 real capital stock per capita, in millions of euro, estimated at the NUTS 2 regional level by Derbyshire et al. (2013), i.e., *Real-Capital*. By including this proxy, I also deal with the possibility that more politically autonomous regions have received larger transfer payments from the central government (Tabellini, 2010). Second, since the Franciscans also heavily affected the medieval rise of the European universities [Knowles 1948, p. 213], I also incorporate in my analysis the variable *Human-Capital*, which is the percentage of the population aged 20-24 enrolled in tertiary education averaged between 2002 and 2009 and available at the NUTS 2 level from Eurostat. Considering *Human-Capital* also reckons with the interplay among human capital, institutions, and growth discussed by a growing literature (Gennaioli et al., 2013). Finally, medieval Western monasticism modulated the intensity of Catholic beliefs influencing in this way present-day economies (see McCleary and Barro, [2006]). Accordingly, I also consider *Catholicism*, which is the share of respondents to the 2008 European Value Study declaring themselves Roman Catholic and answering “very important” to the question “how important is religion in your life?” (GESIS, 2008).

For what finally concerns those dimensions shaping the impact of institutions, I focus on the four factors that have received the closest attention by the related literature (see Olsson and Paik, [2016]). First, Olsson and Paik (2016) claim that, in societies that made an early transition to agriculture in the Neolithic, the persistence of more patriarchal values has delayed the adoption of more inclusive political institutions. To evaluate this aspect, I consider the average time since the agricultural transition in the grid cell calculated exploiting calibrated carbon dates from various Neolithic sites gathered by Pinhasi et al. (2005), i.e., *Neolithic*. Second, Ashraf and Galor (2013) empirically establish that the extent of genetic diversity within a country, as driven by the migratory distance from East Africa, has an

inverted U-shaped relationship with development, a negative effect on generalized trust, and a positive one on conflicts and ethnolinguistic and, in general, cultural fragmentation. Since country-specific genetic diversity is absorbed by the fixed effects, I focus on the Homo Sapiens' exodus out of Africa by including into the specification the migratory distance from Addis Ababa to the centroid of each grid cell, i.e., *Migratory-Distance*.²¹ Third, Iyigun et al. (2015) argue that the soil suitability for potato—i.e., *Potato*—has modulated conflicts and, in turn, institutional evolution in medieval Europe. The raw data are in grid format, cover the entire World at the 0.5° spatial resolution, and were estimated by the GAEZ project. Finally, Voigtländer and Voth (2009) put forward the idea that the Black Death affected both the marriage patterns and the incentive to trade in such a way that the mostly damaged European regions could escape the Malthusian trap. To shed more light on this issue, I control for the mortality rate from the Black Plague in the general population between 1346 and 1353 estimated at the regional level by Benedictow (2004), i.e., *Black Death*.

Turning to the empirical results, table 4 prompts the following observations. First, neither the possibly persistent impact of medieval innovations—i.e., farming progress and Atlantic trades—nor the present-day economic role of the excluded instruments confounds the effect of *Culture* (see columns (3) to (8) of panel A). Crucially, these proxies are not jointly significant in the specifications controlling for all confounding factors (see columns (8) and (9) of panel B). This evidence is consistent with the aforementioned limited relevance of the primary sector, the fact that traveling costs are negligible and thus neither *Ruggedness* nor *Coast* should directly determine outcomes,²² and recent contributions on European regional institutions. In particular, Grafe (2012) documents that in early modern Spain the peripheral regions, home of the autonomous medieval states, obstructed both state formation and market integration to safeguard their own commercial interests losing, in this way, competitiveness over time. Second, *Culture* is not simply picking differences in financial development, human capital, and Christian beliefs driven by the medieval activity of the Cistercians and the Franciscans (see columns (1) to (3) of panel B). Third, none of the factors modulating the

²¹To take into account paleontological and genetic evidence on prehistoric human migration patterns, I always consider Cairo and Istanbul as obligatory intermediate stages (see Ashraf and Galor, [2013]).

²²According to data collected from <http://ec.europa.eu/eurostat>, the average share of household expenditure on transport services (operation of personal transport means) over the sample was about 2 (4) percent.

functioning of permanent institutions modifies the message of my analysis (see columns (4) to (6) of panel B). Finally, conditional on all the confounding factors, the 3SLS estimates are fully consistent with their two-step counterparts and, in particular, a one-standard deviation rise in the strength of a culture of cooperation—i.e., 0.32—will imply an 18% rise in present-day GDP per capita, which is significant at 1%, whereas *Democracy* is again insignificant (see columns (8) and (9) of panel B). For the same regressions moreover, the overidentifying restrictions cannot be rejected at the 75% or more, the Anderson canonical correlations (Sanderson-Windmeijer F) test rejects that equation (1) is underidentified (any endogenous variable is unidentified) at the 7% (8% or less),²³ and geography enters the first-stages in a separable way. To illustrate, *Climate-M* shapes only *Culture*, whereas the direct access to the coast drives only *Democracy*. All in all, these observations suggest that the exclusion restriction holds and the empirical strategy is taking care of reverse causality, the omitted variable bias, and the differential measurement errors in the endogenous regressors.

4.2.2 Semi-Reduced Form Regressions

The validity of the exclusion restriction is also confirmed by the semi-reduced form regressions in table 5. Here, I explicitly address the concern that the excluded instruments might directly affect the economy. To evaluate this possibility, I include one at the time each instrument in both the first- and second-stages. Then, *Culture* has about the same estimated effect as in column (2) of panel A of table 4, and it is always significant at 1%, whereas none among *Democracy*, *Climate-M*, *Ruggedness*, and *Coast* has a significant direct impact on outcomes. Crucially, both the Anderson canonical correlations and the Sanderson-Windmeijer F test reject underidentification except in the case of column (2) where *Ruggedness* is included in the second-stage and thus the first-stage for *Democracy* becomes weak. By “horse racing” predicted institutions with excluded instruments, these regressions stress again that there is no direct significant influence of medieval geography on present-day performance.

4.2.3 Pairwise Analysis of Adjacent Grid Cells

In spite of employing a rich conditioning set, one may still be worried that unobservables are driving the results. To further tackle this issue, I focus on contiguous grid cells with

²³With multiple endogenous regressors, it makes little sense to judge identification from the size of the F-test since each instrument is called upon to play a role in each first-stage (Sanderson and Windmeijer, 2016).

different medieval climate volatility to confirm that the link between culture and development survives even conditional on all unobserved features specific to the relevant 120 km \times 240 km dyads (see also Michalopoulos and Papaioannou, [2013]). This exercise cannot be tailored to contrast culture and inclusive political institutions, but it is naturally fitted to confirm the casual impact of culture. First, I identify contiguous grid cells falling in the same country whose difference in *Climate-M* is at least 0.01 Celsius.²⁴ Next, to avoid that the results are driven by redistribution toward the country administrative center or by pairs with very diverse land area, I exclude the grid cells to which the national capitals belong and those with a land area lower than 200 square km. Finally, I run the second-stage regression

$$Y_{i(j),c} = \alpha_{i(j),c} + \beta_1 C_{i,c} + \delta'_3 \mathbf{X}_{i,c} + \epsilon_{i(j),c}, \quad (3)$$

where $Y_{i(j),c}$ is *Income* in grid cell i of country c that is adjacent to grid cell j of the same country c with grid cells i and j differing in their *Climate-M* values. There are 157 pairs of such grid cells. Since I am now including country-specific, grid cell-pair fixed effects $\alpha_{i(j),c}$, β_1 captures whether differences in medieval climate volatility translate into differences in culture and in turn GDP per capita within pairs of contiguous grid cells in the same country conditional on the rich set of observables contained in $\mathbf{X}_{i,c}$ and unobserved grid cell-pair specific features like local natural resources, technological inputs, and persistent beliefs.

Table 6 reports the results of the contiguous grid cell analysis. The key observations are that *Climate-M* is always a strong predictor of *Culture* and *Income* is significantly higher in the grid cells that display stronger norms of respect and trust today because they experienced a more erratic climate during the Middle Ages, conditional on all observables and unobserved grid cell-pair specific features (see columns (8) and (9) of panel B). In particular, the 3SLS estimates imply that a one-standard deviation rise in *Culture*—i.e., 0.27—will lead to a 10% increase in *Income* and that the overidentifying restrictions cannot be rejected at the 42%.

4.2.4 Falsification Test

Consistent with the first- and second-stages results reported in tables 4 to 6, there is

²⁴This is the first quartile of the strictly positive differences in *Climate-M* between contiguous grid cells. The gist of this section will be the same should I use as threshold either the second or the third quartile.

a positive and significant link between medieval climate volatility and present-day income. To illustrate, the estimated OLS coefficient equals 1.42 with a t-statistic of 9.19 for the sample used in column (2) of table 4 (see left graph in figure 4). Populations that were more exposed to the risk of harvest destruction accumulated a stronger culture of cooperation, and today their descendants are more cooperative and richer. My identification strategy rests on the assumption that risk-sharing-driven cultural accumulation is the only channel through which medieval climate volatility affects current outcomes. If this is true, then a positive relationship between the volatility of the medieval growing season temperature and present-day income should not exist where the cost of accumulating culture was prohibitive. This was the case in Turkey where first the 1058 East-West Schism and then the rise of the Ottoman empire blocked both the Cistercians' and the Franciscans' penetration.²⁵ While indeed the Eastern Orthodox church required that monks shied away from any involvement with the worshipers' life [Tobin 1995, p. 144], Islam considers monasticism an excessive austere practice that thus should be discouraged (The Qur'an, 57.27). I test whether there is no link between medieval climate volatility and present-day economic outcomes in Turkey as follows. First, I divide its surface into 117 grid cells of 1° width. Then, I construct for this sample the variable *Climate-M* and the natural logarithm of the 2009 GDP per capita from the same sources used above. Finally, I condition both variables on the latitude and the longitude of the centroid of the grid cell. As the right graph of figure 4 reveals, there is a negative and insignificant relationship between medieval climate volatility and present-day income in Turkey with an estimated OLS coefficient of - 1.64 and a t-statistic of - 1.52.

5 Inside the Black Box

All in all, it is fair to take stock of the evidence presented so far as consistent with, if not proving, causality going from medieval geography to present-day institutions and the primacy of a culture of cooperation. While an exhaustive account of the mechanisms underlying this last result is beyond the scope of the present paper, in this section I exploit data on the misbehaviors of the members of the House of Representatives of the Italian Parliament

²⁵Our sources report only one (six) Cistercian (Franciscan) house(s)—i.e., Istanbul (Beyoğlu, Istanbul, Izmir, Samsun, Sinop, and Trabzon)—active in Turkey over the 1000-1600 period.

gathered by Chang et al. (2010) to test the idea that culture but not inclusive political institutions is necessary to produce public-spirited politicians and push voters to punish political malfeasance. Ideally, this test would need data on the misbehaviors of all regional representatives in the sample. Yet, it is extremely hard to identify comparable measures of misbehaviors across NUTS 2 regions. Focusing instead on Italian Parliament members has several major advantages. First, a homogeneous measure of political malfeasance is available. Second, autonomous regions are typically run by region-specific parties, which usually obtain the majority also at national elections, and thus more inclusive regional institutions should strengthen the voters' incentive to monitor all their representatives and not only the regional ones. For instance, since 1945, the Südtiroler Volkspartei has represented the interests of Ladin minorities and has gained about two-thirds of the preferences in both the regional and national elections held in the province of Bolzano. Finally, Italian regions exhibit large dissimilarities in culture, the inclusiveness of political institutions, and geography both across and within the Northern and Southern clusters (see figures 1, 2, and 5).

I rely on data from the first to ninth legislatures elected between 1948, year of the first parliamentary election of the Italian Republic, and 1987, last year in which the members of the Parliament enjoyed immunity from criminal prosecution, for 31 of the 32 electoral districts existing at the time. Data for the 31st district of Sardinia are unavailable. Typically these districts group several NUTS 3 Italian units, i.e., *province*. After having dropped politicians with missing values, the total number of observations is 5,755. Immunity could be waived by a vote of Parliament, at the request of the prosecutor. The prosecutor's request to continue with her/his criminal investigation—i.e., *Richiesta di Autorizzazione a Procedere* or RAP from here on—typically received a lot of attention from the media (Nannicini et al., 2013). Accordingly, I focus on a binary turning on whenever the politician received a request by the prosecutor for removal of parliamentary immunity because suspected of a crime, i.e., *RAP*.²⁶ By definition, a RAP is an allegation of malfeasance, rather than a conviction, and as such it could also capture judicial zeal and/or prejudice. Nevertheless, members

²⁶Following the scandals that destroyed the major political parties, the XI legislative term opened the so-called *Second Republic*. Nannicini et al. (2013) also present two measures of political misbehaviors for this period, i.e., the absenteeism rate and the politician's propensity to propose laws targeted to local constituencies. I do not consider these two conducts because much less disruptive and publicized than those eliciting a RAP.

of Parliament could receive a *RAP* from any Italian tribunal and at the provincial level *RAP* is strongly correlated with a measure of corruption based on the extent of missing infrastructures in public works in the 1990s (see Chang et al., [2010]).

Nannicini et al. (2013) propose a model implying that a larger fraction of civic voters discourages moral hazard by politicians. Moreover, a stronger culture of cooperation produces representatives who are less opportunistic and more likely to internalize social welfare. Finally, immoral politicians might self select in low culture districts in search of a lenient electorate. A more inclusive political process, instead, can facilitate the monitoring of politicians by voters but is irrelevant if the latter are not morally compelled to punish political malfeasance or if politicians are inconsiderate (Boix and Posner, 1998; Padró i Miquel et al., 2015). All in all, only culture should be significantly and negatively related to *RAP*.

5.1 Empirical Strategy and Main Results

A glance at figures 1 and 5 confirms the idea just discussed whereby representatives elected in more respectful districts seem more likely to receive a *RAP*, whereas those elected in autonomous regions do not. Next, I verify this remark through multivariate analysis. I add *Democracy* to the Nannicini et al.’s (2013) model, and so I run the second-stage

$$M_{p,d,t} = \kappa_t + \beta_2 C_d + \gamma_1 D_d + \mathbf{X}'_d \delta_4 + \mathbf{Z}'_{p,d,t} \chi + \xi_{p,d,t}, \quad (4)$$

where $M_{p,d,t}$ is *RAP* for politician p , elected in the electoral district d , in the legislature t .²⁷ The excluded instruments for C_d and D_d are *Climate-M*, *Ruggedness*, and *Coast*. The legislature fixed effects κ_t take into account aggregate legislative term shocks, whereas the vector $\mathbf{Z}_{p,d,t}$ gathers individual characteristics like age, education, political experience, and region of birth dummies.²⁸ Finally, \mathbf{X}_d pools the other control variables discussed above except *Latitude* and *Longitude* to avoid collinearity with the region of birth dummies. To

²⁷Switching to an instrumental variables Probit estimator is not feasible since the routine maximizing the relative likelihood function often fails to converge.

²⁸To be precise, $\mathbf{Z}_{p,d,t}$ gathers the member of Parliament’s years of schooling, tenure in legislative terms, age and age squared in years, whether she/he was a minister or vice-minister, whether she/he had previous government experience at the local level, whether her/his previous parliamentary tenure was zero, whether she/he was part of the government coalition, job dummies—i.e., entrepreneur, executive, lawyer, politician, and teacher, legislative term dummies, and region of birth dummies (see for details Nannicini et al., [2013]).

match data measured at the NUTS 2 unit (grid cell) level to districts, I construct averages weighted by each represented unit (grid cell) relative contribution to the district land area.

The estimates in table 7 reveal that the incidence of RAP is significantly lower in districts in which internalized norms of respect and trust are stronger but not in those characterized by more inclusive political institutions. Conditional on all observables indeed, an increase in *Culture* equal to its standard deviation—i.e., 0.23—will reduce the incidence of RAP by about 27 percent, and moving from the lowest level of culture, which is - 0.28 in Puglia, to its highest level, which is 0.92 in Trentino-Alto Adige, will decrease the expected value of *RAP* by about the 139 percent (see columns (7) and (8) of panel B).²⁹ Once again, the consistency of the estimates is confirmed by the underidentification tests and the Sargan statistic.

All in all, I interpret these results as supporting the idea that a culture of cooperation but not more inclusive political institutions significantly strengthens political accountability. Since this constitutes a key instrument through which society can curb the risk of expropriation by politically powerful elites and assure that taxation is properly transformed in public goods, the estimates in table 7 suggest a key mechanism inducing the primacy of culture.

6 Concluding Comments

This paper has exploited exogenous variation created at the European regional level by medieval history to identify the separate roles of present-day culture of cooperation and inclusive political institutions. First, I divide Europe into 120 km \times 120 km grid cells, and I proxy culture with self-reported norms of respect and trust for others and the inclusiveness of the political process with a measure of regional political autonomy. Next, I document strong and distinct first-stage relationships between present-day culture and the severity of consumption risk—i.e., climate volatility—over the 1000-1600 period and between the inclusiveness of present-day political institutions and the factors that raised the returns on elite-citizenry investments in the Middle Ages, i.e., the terrain ruggedness and the direct access to the coast. Building on these first-stages, I report 2SLS estimates suggesting that only culture has a major impact on development even after controlling for country fixed

²⁹A concern with these estimates is that culture discourages criminal prosecution through the behaviors of the judiciary, rather than those of voters. As underlined by Nannicini et al. (2013), this is not very likely since the presence of more zealous judges in high-culture districts might actually increase the likelihood of RAPs.

effects, medieval innovations, the present-day role of medieval geography, intermediate outcomes, and factors modulating the impact of permanent institutions. Crucially, the excluded instruments have no direct impact on development, and the effect of culture holds within pairs of adjacent grid cells with different medieval climate volatility.

To identify a possible channel of causality, I test the idea that more inclusive political institutions are irrelevant in facilitating the monitoring of politicians by voters if the latter are not morally compelled to punish political malfeasance or if the former have weak civic virtues. In particular, I show that there are considerably fewer criminal prosecutions of Italian Parliament members in electoral districts in which culture is stronger but not in districts endowed with more inclusive political institutions. This evidence points at a key mechanism inducing the primacy of culture. Yet, more work is needed to fully characterize the different conduits through which (in)formal institutions affect the economy.

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Tables and Figures

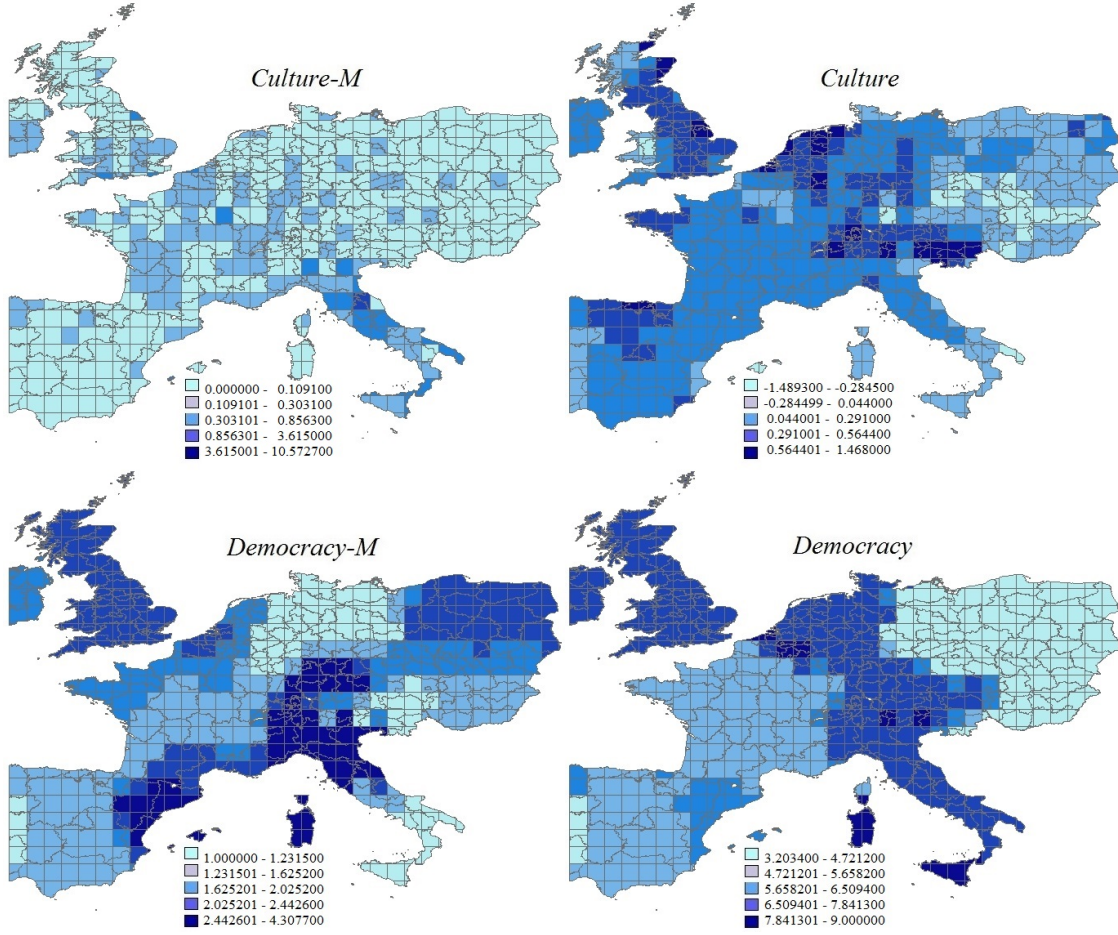
Table 1: The Sample — Medieval States, Historical Regions, and Present-day Countries

| | |
|--|--|
| GENOA: Italy (<i>Liguria</i>); France (<i>Corse</i>). HOLY ROMAN EMPIRE: Austria and Italy (<i>Styria-Austria</i> , <i>Tyrol</i> - <i>Trentino-Alto Adige</i>); Belgium (<i>Région Bruxelles</i> , <i>Région Wallone</i>); Germany (<i>Baden-Württemberg</i> , <i>Bayern</i> , <i>Brandenburg</i> , <i>Bremen</i> - <i>Hamburg</i> - <i>Niedersachsen</i> , <i>Hessen</i> , <i>Mecklenburg-Vorpommern</i> , <i>Nordrhein-Westfalen</i> , <i>Rheinland-Pfalz</i> - <i>Saarland</i> , <i>Sachsen</i> , <i>Schleswig-Holstein</i> , <i>Thüringen</i> - <i>Sachsen-Anhalt</i>); Slovenia (<i>Carniola</i> , <i>Styria-Slovenia</i>). KINGDOM OF BOHEMIA: Czech Republic (<i>East Czech Republic</i> , <i>West Czech Republic</i>); Poland (<i>South Poland</i> , <i>West Poland</i>). KINGDOM OF PORTUGAL: Portugal (<i>Alentejo</i> , <i>Algarve</i> , <i>Centro</i> , <i>Lisboa</i> - <i>Vale do Tejo</i> , <i>Norte</i>). KINGDOM OF SICILY: Italy (<i>Abruzzo</i> - <i>Molise</i> , <i>Basilicata</i> - <i>Campania</i> , <i>Calabria</i> , <i>Puglia</i> , <i>Sicilia</i>). KINGDOM OF TUSCANY: Italy (<i>Toscana</i>). PAPAL STATE: Italy (<i>Emilia-Romagna</i> , <i>Lazio</i> , <i>Marche</i> - <i>Umbria</i>). PROVINCES: Netherlands (<i>Noord Nederland</i> - <i>Groningen</i> , <i>Oost-Nederland</i> , <i>West-Nederland</i> , <i>Zuid-Nederland</i>). REIGN OF ENGLAND: Ireland (<i>East Ireland</i> , <i>West Ireland</i>); UK (<i>East Anglia</i> - <i>London</i> , <i>East Midlands</i> , <i>North-East UK</i> , <i>North-West UK</i> , <i>Northern Ireland</i> , <i>Scotland</i> , <i>South-East UK</i> , <i>South-West UK</i> , <i>Wales</i> , <i>West Midlands</i> , <i>Yorkshire and the Humber</i>). REIGN OF FRANCE: Belgium (<i>Vlaams Gewest</i>); France (<i>East France</i> , <i>Île de France</i> , <i>Mediterranean France</i> , <i>North France</i> , <i>Paris Basin</i> , <i>South-East France</i> , <i>South-West France</i> , <i>West France</i>). REIGN OF HUNGARY: Hungary (<i>Central Hungary</i> , <i>Styria-Hungary</i> , <i>West Hungary</i>); Slovakia (<i>East Slovakia</i> , <i>West Slovakia</i>). REIGN OF POLAND: Poland (<i>East Poland</i> , <i>North Poland</i>). REIGN OF SPAIN: Spain (<i>Andalucía</i> , <i>Aragon</i> , <i>Asturias</i> - <i>Cantabria</i> , <i>Baleares</i> , <i>Castilla-La Mancha</i> , <i>Castilla y León</i> , <i>Cataluña</i> , <i>Comunidad Valenciana</i> , <i>Extremadura</i> , <i>Galicia</i> , <i>Madrid</i> , <i>Murcia</i> , <i>Navarra</i> - <i>Rioja</i> , <i>Pais Vasco</i>). SARDINIAN GIUDICATI: Italy (<i>Sardegna</i>). SAVOY: Italy (<i>Piemonte</i> - <i>Valle D'Aosta</i>). STATE OF MILAN: Italy (<i>Lombardia</i>). SWISS CANTONS: Switzerland (<i>North Switzerland</i> , <i>South Switzerland</i>). VENICE: Italy (<i>Friuli-Venezia Giulia</i> - <i>Veneto</i>). | |
| Note: | 1. The names of the medieval states are in capital font, those of the historical regions that constitute the cross-section identifiers are in italic lowercase type, and those of the present-day countries to which these regions belong are in regular lowercase font. |

Table 2: Summary of Variables

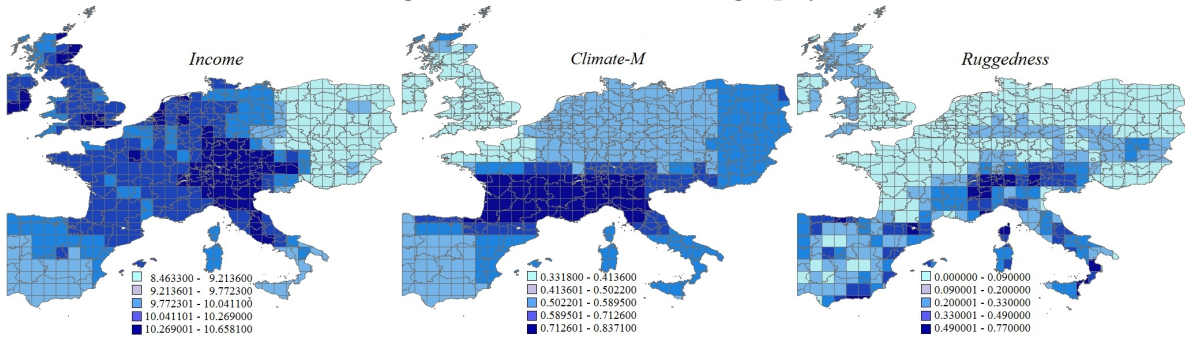
| | Variable | Definition and Sources | Statistics |
|--|---|--|-----------------------|
| Economic outcomes: | <i>Income:</i> | Natural logarithm of the annual GDP per capita in euro averaged over the NUTS 2 regions to which the grid cell belongs and the 2002-2009 period. Source: http://epp.eurostat.ec.europa.eu/ | 9.838 (0.546) |
| | <i>Culture-M:</i> | See text. Sources: http://www.cistercensi.info/ ; Van Der Meer (1965); http://users.bart.nl/~roestb/franciscan ; Moorman (1983). | 0.204 (1.217) |
| Institutions: | <i>Culture:</i> | See text. Source: European Value Study, GESIS (2008). | 0.129 (0.360) |
| | <i>Democracy-M:</i> | Constraints on the elite's power score averaged over the historical regions to which the grid cell belongs and the 1000-1600 period. Source: Boranbay and Guerriero (2016). | 1.816 (0.594) |
| | <i>Democracy:</i> | See text. Sources: Author's codification and Polity IV data set, available at http://www.systemicpeace.org | 5.982 (1.426) |
| Political accountability: | <i>RAP:</i> | Dummy equal to one if the Parliament received a request for removal of the politician's immunity because suspected of a crime. Source: Chang et al. (2010). | 0.233 (0.423) |
| | <i>Climate-M:</i> | Standard deviation of the 1000-1600 growing season temperature in degree Celsius averaged over the grid cells used in Guiot et al. (2010). Source: Guiot et al. (2010). | 0.531 (0.128) |
| Excluded Instruments: | <i>Ruggedness:</i> | Terrain ruggedness in km. Source: http://gecon.yale.edu/ | 0.163 (0.150) |
| | <i>Coast:</i> | Dummy equal to one if the grid cell has a direct access to the Mediterranean or the Atlantic Ocean, 0 otherwise. | 0.367 (0.482) |
| | <i>Latitude:</i> | Latitude of the centroid of the grid cell. | 47.663 (5.348) |
| Other controls: | <i>Longitude:</i> | Longitude of the centroid of the grid cell. | 6.986 (8.802) |
| Other economic roles of excluded instruments: | <i>Primary-Sector:</i> | Share of the active population employed in agriculture and fishing averaged over the NUTS 2 regions to which the grid cell belongs and the 2002-2008 period. Source: http://epp.eurostat.ec.europa.eu/portal/page/portal/statistics/ | 0.065 (0.054) |
| | <i>Atlantic-Trade:</i> | Number of Atlantic ports in the grid cell between 1500 and 1850. Source: Acemoglu et al. (2005). | 0.123 (0.457) |
| | <i>LPD-M:</i> | Natural logarithm of the population per square km averaged over the grids used in Goldewijk et al. (2011) and the 1000-1600 period. Source: Goldewijk et al. (2011). | 3.027 (0.933) |
| | <i>Climate:</i> | Normalized—to range between 0 and 1—first principal component extracted from the standard deviation of the temperature in degree Celsius and that of the precipitation in mm both averaged between 1961 and 1990. Source: http://gecon.yale.edu/ | 0.295 (0.184) |
| | <i>Distance-to-Coast:</i> | Average distance to the coast within the grid cell in km. Source: http://gecon.yale.edu/ | 160.470 (158.693) |
| | <i>Traveling-Distance:</i> | Average traveling distance between the centroid and the corners of the grid cell in km. Source: http://www.distancefromto.net | 256.548 (289.872) |
| | <i>Real-Capital:</i> | Real capital stock per capita in 2000, in millions of euro, averaged over the NUTS 2 regions to which the grid cell belongs. Source: Derbyshire et al. (2013). | 0.052 (0.026) |
| Intermediate Outcomes: | <i>Human-Capital:</i> | Percentage of the population aged 20-24 enrolled in tertiary education—i.e., ISCED 5-6—averaged over the NUTS 2 regions to which the grid cell belongs and the 2002-2009 period. Source: http://epp.eurostat.ec.europa.eu/ | 51.057 (15.845) |
| | <i>Catholicism:</i> | See text. Source: European Value Study, GESIS (2008). | 0.258 (0.152) |
| | <i>Neolithic:</i> | Average time since agricultural transition in years. Source: Pinhasi et al. (2005). | 6747.321 (657.154) |
| Dimensions shaping the impact of institutions: | <i>Migratory-Distance:</i> | Migratory distance from Addis Ababa to the grid cell centroid in thousands of km. Source: Ashraf and Galor (2013). | 5.704 (0.626) |
| | <i>Potato:</i> | Land suitability for white potato ranging between 0 and 100 and averaged over the grids used in the GAEZ data set. Source: http://www.gaez.iiasa.ac.at/ | 26.335 (13.756) |
| | <i>Black-Death:</i> | Mortality rate from Black Plague in the population between 1346 and 1353 averaged over the NUTS2 regions to which the grid cell belongs. Source: Benedictow (2004). | 59.542 (2.999) |
| Note: | 1. The last column reports the mean and, in parentheses, the standard deviation of each variable. Both are computed building on the samples used in tables 3 and 4 except in the case of <i>RAP</i> , when they are calculated exploiting the sample used in table 7. | | |

Figure 1: Persistent Institutions



Note: 1. The range of each variable is divided into five intervals using the goodness of variance fit method.

Figure 2: Income and Geography



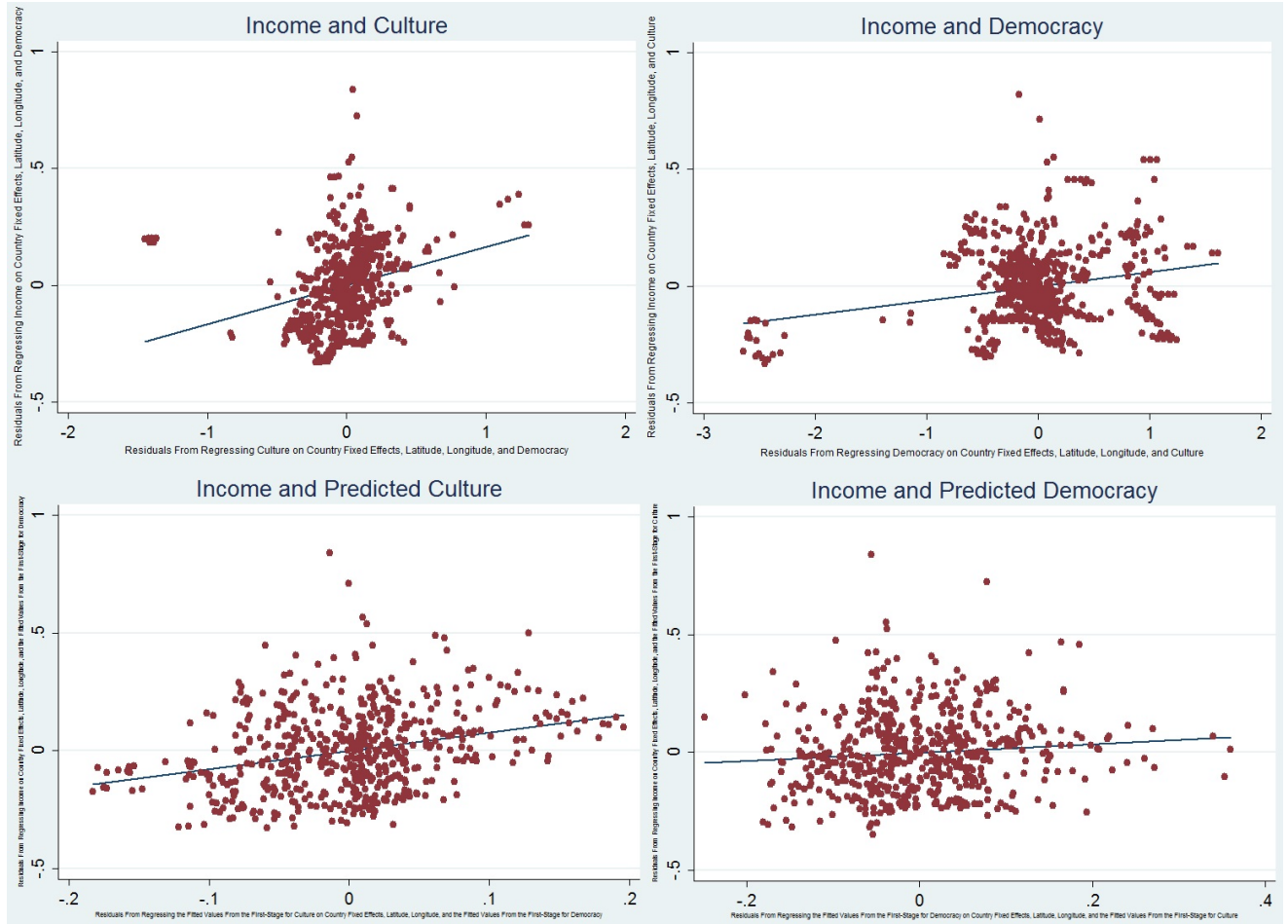
Note: 1. The range of each variable is divided into five intervals using the goodness of variance fit method.

Table 3: Persistent Endogenous Institutions

| | (1) | (2) | The dependent variable is: | |
|--|-----------------------|-----------------------|----------------------------|---------------------|
| | <i>Culture</i> | <i>Culture</i> | <i>Democracy</i> | <i>Democracy</i> |
| <i>Culture-M</i> | 0.046 (0.009)*** | | 0.137 (0.055)*** | |
| <i>Democracy-M</i> | | | | |
| <i>Climate-M</i> | | 0.398 (0.125)*** | 0.017 (0.286) | - 0.025 (0.294) |
| <i>Ruggedness</i> | 0.255 (0.097)*** | 0.155 (0.103) | | 0.688 (0.242)*** |
| <i>Coast</i> | - 0.111 (0.028)*** | - 0.101 (0.028)*** | | 0.112 (0.067)* |
| p-value for <i>Latitude</i> and <i>Longitude</i> | [0.00] | [0.00] | [0.00] | [0.00] |
| Estimation | OLS | | | |
| Within R ² | 0.17 | 0.15 | 0.07 | 0.08 |
| Number of observations | 578 | 578 | 578 | 578 |

Notes: 1. Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
2. All specifications also consider country fixed effects.

Figure 3: Institutions and Outcomes — OLS Versus 2SLS



Note: 1. The upper-left(right) graph depicts the significant increasing effect, estimated using OLS, of *Culture* (*Democracy*) on *Income* in the 578-grid cell sample used in column (1) of panel A of table 4, conditional on *Latitude*, *Longitude*, and country fixed effects. The bottom-left(right) graph depicts the significant (insignificant) increasing effect, estimated using 2SLS with excluded instruments *Climate-M*, *Ruggedness*, and *Coast*, of *Culture* (*Democracy*) on *Income* in the 578-grid cell sample used in column (2) of panel A of table 4, conditional on *Latitude*, *Longitude*, and country fixed effects.

Table 4: Institutions and Outcomes — Country Fixed Effects OLS, 2SLS, and 3SLS

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|---|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|---------------------|
| Panel A. The dependent variable is <i>Income</i> | | | | | | | | |
| <i>Culture</i> | 0.165 (0.026)*** | 0.779 (0.158)*** | 0.778 (0.156)*** | 0.765 (0.146)*** | 0.874 (0.183)*** | 0.641 (0.430) | 0.923 (0.200)*** | 0.849 (0.282)*** |
| <i>Democracy</i> | 0.061 (0.011)*** | 0.176 (0.116) | 0.175 (0.114) | 0.180 (0.110)* | 0.278 (0.144)** | 0.380 (0.521) | 0.094 (0.109) | 0.138 (0.165) |
| <i>Primary-Sector</i> | | | - 0.038 (0.205) | | | | | |
| <i>Atlantic-Trade</i> | | | | - 0.009 (0.028) | | | | |
| <i>LPD-M</i> | | | | | 0.055 (0.016)*** | | | |
| <i>Climate</i> | | | | | | - 0.356 (0.712) | | |
| <i>Distance-to-Coast</i> | | | | | | | - 0.0002 (0.0001)* | |
| <i>Traveling-Distance</i> | | | | | | | | 0.00003 (0.0001) |
| First-Stage for <i>Culture</i> | | | | | | | | |
| <i>Climate-M</i> | | 0.398 (0.125)*** | 0.418 (0.128)*** | 0.422 (0.124)*** | 0.398 (0.125)*** | 0.317 (0.123)*** | 0.407 (0.125)*** | 0.373 (0.124)*** |
| <i>Ruggedness</i> | | 0.155 (0.103) | 0.152 (0.104) | 0.168 (0.102)* | 0.156 (0.105) | 0.332 (0.106)*** | 0.151 (0.103) | 0.152 (0.102) |
| <i>Coast</i> | | - 0.101 (0.028)*** | - 0.103 (0.029)*** | - 0.116 (0.029)*** | - 0.101 (0.028)*** | - 0.039 (0.030) | - 0.076 (0.035)** | - 0.043 (0.035) |
| Sanderson-Windmeijer test p-value | | 0.00 | 0.00 | 0.00 | 0.00 | 0.66 | 0.00 | 0.04 |
| First-Stage for <i>Democracy</i> | | | | | | | | |
| <i>Climate-M</i> | | - 0.025 (0.294) | - 0.032 (0.300) | - 0.062 (0.294) | - 0.047 (0.295) | 0.103 (0.293) | 0.012 (0.293) | 0.015 (0.294) |
| <i>Ruggedness</i> | | 0.688 (0.242)*** | 0.694 (0.244)*** | 0.668 (0.242)*** | 0.640 (0.247)*** | 0.407 (0.253) | 0.676 (0.242)*** | 0.692 (0.242)*** |
| <i>Coast</i> | | 0.112 (0.067)* | 0.118 (0.067)* | 0.135 (0.068)** | 0.111 (0.067)* | 0.014 (0.072) | 0.212 (0.082)*** | 0.021 (0.082) |
| Sanderson-Windmeijer test p-value | | 0.00 | 0.00 | 0.00 | 0.01 | 0.71 | 0.00 | 0.09 |
| Estimation | OLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS |
| Within R ² | 0.15 | | | | | | | |
| P-value of underidentification test | | 0.00 | 0.00 | 0.00 | 0.01 | 0.71 | 0.00 | 0.10 |
| P-value of Sargan statistic | | 0.57 | 0.58 | 0.55 | 0.59 | 0.54 | 0.85 | 0.62 |
| Number of observations | 578 | 578 | 573 | 578 | 578 | 578 | 578 | 578 |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Panel B. The dependent variable is <i>Income</i> | | | | | | | | |
| <i>Culture</i> | 0.748 (0.303)*** | 0.734 (0.165)*** | 0.822 (0.165)*** | 0.907 (0.198)*** | 0.878 (0.193)*** | 0.771 (0.161)*** | 0.062 (0.025)*** | 0.568 (0.192)*** |
| <i>Democracy</i> | 0.206 (0.140) | 0.159 (0.110) | 0.181 (0.126) | - 0.066 (0.115) | 0.277 (0.178) | 0.205 (0.132) | 0.006 (0.009) | 0.044 (0.114) |
| <i>Real-Capital</i> | 0.702 (3.401) | | | | | | | |
| <i>Human-Capital</i> | | 0.002 (0.001)* | | | | | | |
| <i>Catholicism</i> | | | 0.255 (0.189) | | | | | |
| <i>Neolithic</i> | | | | - 6E-6 (0.00003) | | | | |
| <i>Migratory-Distance</i> | | | | - 0.892 (0.152)*** | | | | |
| <i>Potato</i> | | | | | 0.002 (0.002) | | | |
| <i>Black-Death</i> | | | | | | - 0.010 (0.007) | | |
| P-value for medieval innovations and current role of medieval geography | | | | | | | [0.00] | [0.18] |
| P-value for all extra controls | | | | | | | [0.00] | [0.00] |
| First-Stage for <i>Culture</i> | | | | | | | | |
| <i>Climate-M</i> | 0.359 (0.125)*** | 0.293 (0.124)** | 0.395 (0.125)*** | 0.391 (0.107)*** | 0.421 (0.129)*** | 0.397 (0.125)*** | 0.277 (0.109)*** | 0.295 (0.090)*** |
| <i>Ruggedness</i> | 0.043 (0.107) | 0.145 (0.101) | 0.154 (0.103) | 0.202 (0.092)** | 0.193 (0.118)* | 0.154 (0.103) | 0.276 (0.107)*** | 0.260 (0.090)*** |
| <i>Coast</i> | - 0.077 (0.028)*** | - 0.107 (0.028)*** | - 0.101 (0.028)*** | - 0.031 (0.026) | - 0.099 (0.028)*** | - 0.101 (0.029)*** | - 0.003 (0.033) | - 0.002 (0.031) |
| Sanderson-Windmeijer test p-value | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | |
| First-Stage for <i>Democracy</i> | | | | | | | | |
| <i>Climate-M</i> | - 0.335 (0.272) | - 0.013 (0.298) | - 0.174 (0.285) | - 0.071 (0.314) | - 0.072 (0.304) | - 0.068 (0.293) | - 0.214 (0.306) | - 0.211 (0.295) |
| <i>Ruggedness</i> | 0.393 (0.234)* | 0.690 (0.243)*** | 0.647 (0.234)*** | 0.738 (0.272)*** | 0.566 (0.278)** | 0.647 (0.242)*** | 0.051 (0.301) | 0.048 (0.290) |
| <i>Coast</i> | 0.237 (0.062)*** | 0.113 (0.067)* | 0.118 (0.064)* | 0.118 (0.077) | 0.111 (0.067)* | 0.089 (0.067) | 0.203 (0.092)** | 0.203 (0.089)** |
| Sanderson-Windmeijer test p-value | 0.02 | 0.00 | 0.00 | 0.02 | 0.04 | 0.01 | 0.08 | |
| Estimation | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | OLS | 2SLS |
| Within R ² | | | | | | | 0.67 | 3SLS |
| P-value of underidentification test | 0.03 | 0.00 | 0.00 | 0.02 | 0.04 | 0.01 | 0.07 | |
| P-value of Sargan statistic | 0.57 | 0.78 | 0.49 | 0.79 | 0.59 | 0.44 | 0.75 | 0.95 |
| Number of observations | 563 | 578 | 578 | 518 | 577 | 578 | 500 | 500 |

- Notes:
- Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
 - All specifications also consider *Latitude*, *Longitude*, and country fixed effects. The extra controls included in the specifications in columns (7) to (9) of panel B are *Primary-Sector*, *Atlantic-Trade*, *LPD-M*, *Climate*, *Distance-to-Coast*, *Traveling-Distance*, *Real-Capital*, *Human-Capital*, *Catholicism*, *Neolithic*, *Migratory-Distance*, *Potato*, and *Black-Death*.
 - In columns (2) to (8) of panel A and columns (1) to (6) and (8) of panel B (column (9) of panel B), the endogenous variables are *Culture* and *Democracy* (*Income*, *Culture*, and *Democracy*) and the excluded instruments are *Climate-M*, *Ruggedness*, and *Coast*. The control variables used in the second-stage are also included in the first-stage.
 - The proxies for medieval innovations are *Primary-Sector*, *Atlantic-Trade*, and *LPD-M*, whereas those for the current role of medieval geography are *Climate*, *Distance-to-Coast*, and *Traveling-Distance*.
 - The null hypothesis of the Sanderson-Windmeijer F test is that the endogenous variable is unidentified, and that of the Anderson underidentification (Sargan) test is that the excluded instruments are uncorrelated with the endogenous variables (exogenous).

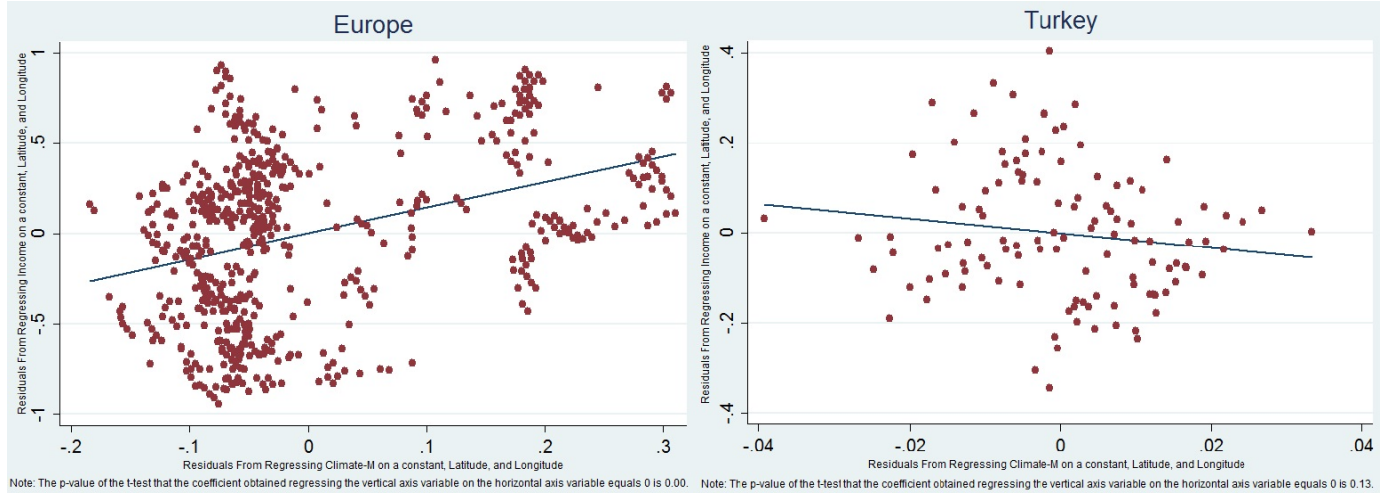
Table 5: Institutions and Outcomes — Semi-reduced Form Regressions

| | (1) | (2) | (3) |
|---|---|---|---------------------|
| | | The dependent variable is <i>Income</i> | |
| <i>Culture</i> | 0.677 (0.219)*** | 0.929 (0.413)** | 0.912 (0.302)*** |
| <i>Democracy</i> | 0.161 (0.109) | 0.387 (0.522) | 0.108 (0.179) |
| <i>Climate-M</i> | 0.095 (0.152) | | |
| <i>Ruggedness</i> | | - 0.195 (0.460) | |
| <i>Coast</i> | | | 0.030 (0.056) |
| Sanderson-Windmeijer test p-value in the first-stage for <i>Culture</i> | 0.00 | 0.24 | 0.02 |
| Sanderson-Windmeijer test p-value in the first-stage for <i>Democracy</i> | 0.00 | 0.33 | 0.04 |
| Estimation | | 2SLS | |
| P-value of underidentification test | 0.00 | 0.32 | 0.04 |
| Number of observations | 578 | 578 | 578 |
| Notes: | <ol style="list-style-type: none"> Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%. All specifications also consider <i>Latitude</i>, <i>Longitude</i>, and country fixed effects. The control variables used in the second-stages are also included in the first-stages, which are as in column (2) of panel A of table 4 and thus not reported in the present table. The endogenous variables are <i>Culture</i> and <i>Democracy</i>, whereas the excluded instruments in columns (1) to (3) are respectively <i>Ruggedness</i> and <i>Coast</i>, <i>Climate-M</i> and <i>Coast</i>, and <i>Climate-M</i> and <i>Ruggedness</i>. The null hypothesis of the Sanderson-Windmeijer F test is that the endogenous variable is unidentified, and that of the Anderson underidentification test is that the excluded instruments are uncorrelated with the endogenous variables. | | |

Table 6: Institutions and Outcomes — Pairwise Analysis of Adjacent Grid Cells

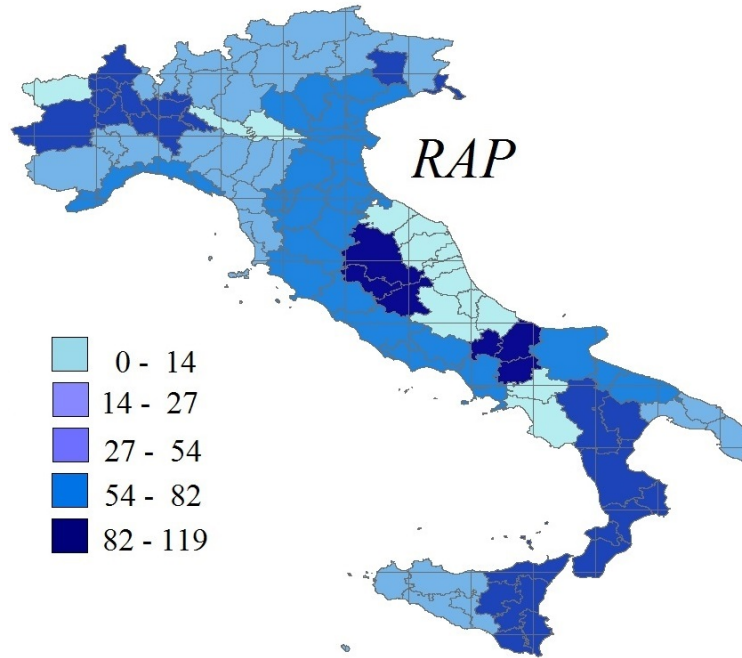
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|---|--|-------------------|-------------------|--------------------|-------------------|------------------|---------------------|---------------------|
| | Panel A. The dependent variable is <i>Income</i> | | | | | | | |
| <i>Culture</i> | 0.199 (0.034)*** | 0.034 (0.288) | 0.013 (0.283) | 0.046 (0.293) | 0.075 (0.259) | 0.042 (0.282) | 0.031 (0.293) | 0.072 (0.294) |
| | First-Stage for <i>Culture</i> | | | | | | | |
| <i>Climate-M</i> | | 0.417 (0.269) | 0.468 (0.290)* | 0.406 (0.270) | 0.439 (0.268)* | 0.424 (0.270) | 0.412 (0.270) | 0.400 (0.272) |
| Sanderson-Windmeijer test p-value | | 0.12 | 0.11 | 0.13 | 0.10 | 0.12 | 0.13 | 0.14 |
| Estimation | OLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS |
| Within R ² | 0.19 | | | | | | | |
| P-value of underidentification test | | 0.12 | 0.11 | 0.13 | 0.10 | 0.11 | 0.13 | 0.14 |
| Number of observations | 314 | 314 | 302 | 314 | 314 | 314 | 314 | 314 |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| | Panel B. The dependent variable is <i>Income</i> | | | | | | | |
| <i>Culture</i> | 0.175 (0.171) | 0.135 (0.217) | 0.062 (0.298) | - 0.306 (0.780) | 0.064 (0.253) | 0.033 (0.301) | 0.183 (0.042)*** | 0.371 (0.171)** |
| P-value for medieval innovations and current role of medieval geography | | | | | | | [0.06] | [0.04] |
| P-value for all extra controls | | | | | | | [0.00] | [0.00] |
| | First-Stage for <i>Culture</i> | | | | | | | |
| <i>Climate-M</i> | 0.600 (0.229)*** | 0.449 (0.267)* | 0.395 (0.270) | 0.263 (0.284) | 0.484 (0.281)* | 0.400 (0.264) | | 0.669 (0.253)*** |
| Sanderson-Windmeijer test p-value | 0.01 | 0.09 | 0.15 | 0.36 | 0.09 | 0.13 | | 0.01 |
| Estimation | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | OLS | 2SLS |
| Within R ² | | | | | | | 0.66 | |
| P-value of underidentification test | 0.01 | 0.09 | 0.14 | 0.35 | 0.08 | 0.13 | | 0.01 |
| P-value of Sargan statistic | | | | | | | | 0.42 |
| Number of observations | 314 | 314 | 314 | 256 | 314 | 314 | 246 | 246 |
| Notes: | <ol style="list-style-type: none"> Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%. All specifications also consider <i>Latitude</i>, <i>Longitude</i>, and adjacent grid cells fixed effects. The specifications in columns (3) to (8) of panel A ((1) to (6) of panel B) also include <i>Primary-Sector</i>, <i>Atlantic-Trade</i>, <i>LPD-M</i>, <i>Climate</i>, <i>Distance-to-Coast</i>, and <i>Traveling-Distance</i> (<i>Real-Capital</i>, <i>Human-Capital</i>, <i>Catholicism</i>, <i>Neolithic</i> and <i>Migratory-Distance</i>, <i>Potato</i>, and <i>Black-Death</i>), respectively. The extra controls considered in the specifications reported in columns (7) to (9) of panel B are <i>Primary-Sector</i>, <i>Atlantic-Trade</i>, <i>LPD-M</i>, <i>Climate</i>, <i>Distance-to-Coast</i>, <i>Traveling-Distance</i>, <i>Real-Capital</i>, <i>Human-Capital</i>, <i>Catholicism</i>, <i>Neolithic</i>, <i>Migratory-Distance</i>, <i>Potato</i>, and <i>Black-Death</i>. The control variables used in the second-stage are also included in the first-stage. In columns (2) to (9) of panel A and columns (1) to (6) and (8) of panel B (column (9) of panel B), the endogenous variable(s) is (are) <i>Culture</i> (<i>Income</i> and <i>Culture</i>) and the excluded instrument is <i>Climate-M</i>. The proxies for medieval innovations are <i>Primary-Sector</i>, <i>LPD-M</i>, and <i>Atlantic-Trade</i>, whereas those for the current role of medieval geography are <i>Climate</i>, <i>Traveling-Distance</i>, and <i>Distance-to-Coast</i>. The null hypothesis of the Sanderson-Windmeijer F test is that the endogenous variable is unidentified, and that of the Anderson underidentification (Sargan) test is that the excluded instruments are uncorrelated with the endogenous variables (exogenous). | | | | | | | |

Figure 4: Severity of Consumption Risk and Outcomes — Placebo Test



Note: 1. The left graph depicts the significant increasing effect, estimated using OLS, of *Climate-M* on *Income* in the 578-grid cell sample used in column (1) of panel A of table 4, conditional on *Latitude* and *Longitude*. The right graph depicts the insignificant decreasing effect, estimated using OLS, of *Climate-M* on *Income* in a sample of 117 grid cells covering Turkey, conditional on *Latitude* and *Longitude*.

Figure 5: Malfeasance by the Italian First Republic Parliament



Note: 1. The range of each variable is divided into five intervals using the goodness of variance fit method.

Table 7: Institutions and Political Accountability — The Case of the First Republic in Italy

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|---|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|-----------------------|------------------------|------------------------|
| Panel A. The dependent variable is <i>RAP</i> | | | | | | | | |
| <i>Culture</i> | - 0.104 (0.041)*** | - 0.348 (0.121)*** | - 0.325 (0.139)** | - 0.361 (0.126)*** | - 0.395 (0.133)*** | - 0.428 (0.222)** | - 0.370 (0.112)*** | - 0.354 (0.144)** |
| <i>Democracy</i> | 0.009 (0.015) | - 0.072 (0.088) | - 0.060 (0.075) | - 0.082 (0.091) | - 0.145 (0.104) | - 0.112 (0.143) | - 0.083 (0.084) | - 0.060 (0.084) |
| <i>Primary-Sector</i> | | | 0.132 (0.625) | | | | | |
| <i>LPD-M</i> | | | | - 0.018 (0.024) | | | | |
| <i>Climate</i> | | | | | 0.208 (0.100)** | | | |
| <i>Distance-to-Coast</i> | | | | | | 0.0002 (0.0004) | | |
| <i>Traveling-Distance</i> | | | | | | | - 0.00002 (0.00005) | |
| <i>Real-Capital</i> | | | | | | | | 0.487 (1.341) |
| First-Stage for <i>Culture</i> | | | | | | | | |
| <i>Climate-M</i> | | 2.004 (0.029)*** | 2.088 (0.032)*** | 2.005 (0.028)*** | 2.252 (0.031)*** | 2.075 (0.029)*** | 2.041 (0.030)*** | 1.596 (0.037)*** |
| <i>Ruggedness</i> | | 0.405 (0.016)*** | 0.393 (0.016)*** | 0.344 (0.016)*** | 0.181 (0.019)*** | 0.501 (0.017)*** | 0.398 (0.016)*** | 0.409 (0.015)*** |
| <i>Coast</i> | | - 0.007 (0.004) | - 0.013 (0.004)*** | - 0.016 (0.004)*** | - 0.021 (0.004)*** | - 0.042 (0.005)*** | - 0.013 (0.004)*** | 0.007 (0.004)* |
| Sanderson-Windmeijer test p-value | | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| First-Stage for <i>Democracy</i> | | | | | | | | |
| <i>Climate-M</i> | | - 3.199 (0.110)*** | - 4.334 (0.120)*** | - 3.198 (0.109)*** | - 3.336 (0.121)*** | - 3.407 (0.111)*** | - 3.050 (0.114)*** | - 3.135 (0.143)*** |
| <i>Ruggedness</i> | | - 0.010 (0.060) | 0.153 (0.059)*** | - 0.111 (0.063)* | - 0.114 (0.076) | - 0.296 (0.067)*** | - 0.039 (0.061) | - 0.010 (0.060) |
| <i>Coast</i> | | - 0.131 (0.016)*** | - 0.044 (0.016)*** | - 0.148 (0.016)*** | - 0.123 (0.016)*** | - 0.026 (0.019) | - 0.159 (0.017)*** | - 0.134 (0.016)*** |
| Sanderson-Windmeijer test p-value | | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| Estimation | OLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS |
| Within R^2 | 0.08 | | | | | | | |
| P-value of underidentification test | | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 |
| P-value of Sargan statistic | | 0.68 | 0.72 | 0.91 | 0.24 | 0.53 | 0.55 | 0.54 |
| Number of observations | 5755 | 5755 | 5755 | 5755 | 5755 | 5755 | 5755 | 5755 |
| Panel B. The dependent variable is <i>RAP</i> | | | | | | | | |
| <i>Culture</i> | - 0.344 (0.119)*** | - 0.349 (0.120)*** | - 0.329 (0.120)*** | - 0.391 (0.154)*** | - 0.226 (0.123)* | 0.152 (0.081)* | - 1.161 (0.494)** | - 1.161 (0.494)** |
| <i>Democracy</i> | - 0.071 (0.087) | - 0.081 (0.081) | - 0.035 (0.077) | - 0.103 (0.113) | - 0.003 (0.089) | - 0.001 (0.016) | - 0.009 (0.022) | - 0.009 (0.022) |
| <i>Human-Capital</i> | - 0.0001 (0.0005) | | | | | | | |
| <i>Catholicism</i> | | 0.169 (0.184) | | | | | | |
| <i>Neolithic</i> | | | 1.9E-6 (0.00003) | | | | | |
| <i>Migratory-Distance</i> | | | 0.028 (0.198) | | | | | |
| <i>Potato</i> | | | | - 0.003 (0.002) | | | | |
| <i>Black-Death</i> | | | | | 0.022 (0.009)** | | | |
| P-value for medieval innovations and current role of medieval geography | | | | | | [0.02] | [0.06] | [0.06] |
| P-value for all extra controls | | | | | | [0.00] | [0.07] | [0.07] |
| First-Stage for <i>Culture</i> | | | | | | | | |
| <i>Climate-M</i> | 1.956 (0.027)*** | 2.154 (0.028)*** | 2.240 (0.046)*** | 2.022 (0.029)*** | 2.059 (0.028)*** | | 0.167 (0.056)*** | 0.167 (0.056)*** |
| <i>Ruggedness</i> | 0.542 (0.016)*** | 0.391 (0.015)*** | 0.422 (0.017)*** | 0.327 (0.019)*** | 0.533 (0.017)*** | | 0.290 (0.025)*** | 0.290 (0.025)*** |
| <i>Coast</i> | 0.016 (0.004)*** | - 0.002 (0.004) | - 0.023 (0.004)*** | - 0.010 (0.004)** | - 0.011 (0.004)*** | | - 0.017 (0.005)*** | - 0.017 (0.005)*** |
| Sanderson-Windmeijer test p-value | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | | 0.00 | |
| First-Stage for <i>Democracy</i> | | | | | | | | |
| <i>Climate-M</i> | - 3.156 (0.110)*** | - 3.652 (0.108)*** | - 8.293 (0.140)*** | - 3.115 (0.109)*** | - 3.248 (0.110)*** | | - 16.215 (0.178)*** | - 16.215 (0.177)*** |
| <i>Ruggedness</i> | - 0.131 (0.064)** | 0.034 (0.058) | - 0.595 (0.051)*** | - 0.387 (0.073)*** | - 0.124 (0.068)* | | 0.339 (0.079)*** | 0.340 (0.079)*** |
| <i>Coast</i> | - 0.152 (0.016)*** | - 0.144 (0.015)*** | 0.081 (0.013)*** | - 0.149 (0.016)*** | - 0.128 (0.016)*** | | - 0.106 (0.015)*** | - 0.106 (0.015)*** |
| Sanderson-Windmeijer test p-value | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | | 0.00 | |
| Estimation | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | OLS | 2SLS | 3SLS |
| Within R^2 | | | | | | 0.08 | | |
| P-value of underidentification test | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | | 0.00 | |
| P-value of Sargan statistic | 0.68 | 0.71 | 0.50 | 0.50 | 0.34 | | 0.82 | 0.99 |
| Number of observations | 5755 | 5755 | 5755 | 5755 | 5755 | 5755 | 5755 | 5755 |

- Notes: 1. Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
2. All specifications also consider *Latitude*, *Longitude*, and the regressors listed in footnote 28. The extra controls included in the specifications reported in columns (6) to (8) of panel B are *Primary-Sector*, *LPD-M*, *Climate*, *Distance-to-Coast*, *Traveling-Distance*, *Real-Capital*, *Human-Capital*, *Catholicism*, *Neolithic*, *Migratory-Distance*, *Potato*, and *Black-Death*. The control variables used in the second-stage are also included in the first-stage.
3. In columns (2) to (9) of panel A and columns (1) to (5) and (7) of panel B (column (8) of panel B), the endogenous variables are *Culture* and *Democracy* (*RAP*, *Culture*, and *Democracy*) and the excluded instruments are *Climate-M*, *Ruggedness*, and *Coast*.
4. The proxies for medieval innovations are *Primary-Sector* and *LPD-M*, whereas those for the current role of medieval geography are *Climate*, *Distance-to-Coast*, and *Traveling-Distance*.
5. The null hypothesis of the Sanderson-Windmeijer F test is that the endogenous variable is unidentified, and that of the Anderson underidentification (Sargan) test is that the excluded instruments are uncorrelated with the endogenous variables (exogenous).

APPENDIX (FOR ONLINE PUBLICATION)

Constructing the *Political-Autonomy* Index

Political-Autonomy equals 1 if a region has exclusive control over a limited set of policy—e.g., education, 2 if it is fiscally decentralized, 3 if it has substantial political autonomy from the central government, and 0 otherwise. I consider a region as fiscally decentralized if it can raise part of its fiscal revenues through region-specific taxes and spend them on local public goods. I treat a unit as politically autonomous if it is fiscally decentralized, can elect its own parliament, and controls all policies except those of national relevance like defense. Next, I clarify how I computed the index for the NUTS 2 regions in the sample.

Austria

The 1971 Austro-Italian treaty stipulated that the South Tyrol region would be granted an autonomous status in all similar to the one of the Trentino-Alto Adige within Italy. The autonomy recognized by the special statute covers the political, legislative, administrative, and fiscal institutions with very limited legislative or executive competencies left to the central government (Parolari, 2012). Thus, I assigned a score of 2 to South Tyrol from 1971 on and 1 to the remaining NUTS 2 regions belonging to Austria.

Belgium

Starting from the four language areas (the Dutch, bilingual, French and German language areas), the 1962-3, 1970, and following revisions of the Belgian constitution established that Belgium is a unique federal state with two segregated political power—i.e., region Wallone and the Vlaams Gewest—with independent political and taxation power (Verbeke, 2012). The overlapping boundaries of the Regions and Communities have created two notable peculiarities: the territory of the Brussels-Capital Region is included in both the Flemish and French Communities, and the territory of the German-speaking Community lies wholly within the Walloon Region. Jurisdictional conflicts are resolved by the Constitutional Court.

The Federal State's authority includes justice, defense, federal police, social security, nuclear energy, monetary, fiscal, and foreign policies, and other aspects of public finances which together amount to the 50% of the national fiscal income (Verbeke, 2012). The

communities exercise their authority on policies related to education, use of language, health policy, employment, agriculture, water policy, housing, public works, energy, transport, the environment, town and country planning, nature conservation, and credit. They supervise the provinces, municipalities and inter-communal utility companies. There is almost no possible veto by the Belgian State and, often, Belgium is not even able to sign an international treaty without the agreement of the Walloon and Flemish Parliaments (Verbeke, 2012).

Based on these peculiarities, the Région Wallonne and Vlaams Gewest are assigned a value of 3 for 1962 on whereas the Brussels-Capital region has always a value of 1.

France

With the 1992 Joxe statute, the central government has granted to a Corse Parliament exclusive powers on local policies but without fiscal privileges (Chaubin et al., 2003). I assigned a score of 1 to Corse from 1992 on and 0 to the other French regions.

Italy

Article 116 of the Italian Constitution (1948) grants to the regions of Friuli-Venezia Giulia, Sardegna, Sicilia, Trentino Alto Adige, and Valle D'Aosta not only the power of legislate over public schools, health-care, and local infrastructures but also to retain the vast majority (at least the 70%) of their own tax revenue. With the constitutional decree n. 2 released on the 31/01/2001, the regions with ordinary statutes acquired residual legislative powers. In particular, they now have exclusive legislative power with respect to any matters not expressly reserved to state law (Article 117). Yet their financial autonomy is not complete and they can keep only 20% of all levied taxes, mostly used to finance the region-based healthcare system (Pennino, 2009). On top of this discussion, I have assigned a score of 2 to Friuli-Venezia Giulia, Sardegna, Sicilia, Trentino Alto Adige, and Valle D'Aosta from 1948 on and a score of 1 to all the other regions between 2001 and 2010.

Spain

The starting point in the territorial organization of Spain was the second article of the 1978 constitution, which gave the way to an eventual process of devolution to be realized according to two possible “routes” (Beltrán et al., 2005). The “fast track” was established in article 151, and was implicitly reserved for the three “historical nationalities” of the Cataluna,

Galicia, and Pais Vasco constituted in 1979. To this groups was added Navarra (1982), which acceded to autonomy through the recognition of its historical “charters” and as such it is known as a “chartered community”. The constitution also explicitly established that the institutional framework for these communities would be a parliamentary system, with a Legislative Assembly elected by universal suffrage, a cabinet or “council of government”, a president of such a council, elected by the Assembly, and a Supreme Court of Justice. They were also granted a maximum level of devolved competences. The “slow track” was established in article 143. This route was taken by the other Spanish communities which got constituted in the following years (Andalucia, 1981; Aragon, 1982; Asturias, 1981; Baleares, 1983; Cantabria, 1982; CastillaLa Mancha, 1982; Castilla-Len, 1983; Extremadura, 1983; La Rioja, 1982; Madrid, 1983; Murcia, 1982; Comunidad Valenciana, 1982) and acquired in the 80s and 90s a very similar structure to the one devised for the “historical nationalities”.

Since the late 90s then, all regions have acquired the power to manage their own finances and are responsible for the administration of education, health and social services, and cultural and urban development (Beltrán et al., 2005). Yet, Aragon, Baleares, Cataluna, Comunidad Valenciana, Galicia, and Pais Vasco still keep a wider control of policy-making to the point of adopting a regional civil code (Beltrán et al., 2005). Thus, I assigned starting with the year of foundation of each community a value of 3 to Aragon, Baleares, Cataluna, Comunidad Valenciana, Galicia, and Pais Vasco and 2 to the remaining regions.

United Kingdom

Northern Ireland.—Since the 1998 Good Friday Agreement, Northern Ireland has devolved government within the United Kingdom. The UK Government and UK Parliament are responsible for reserved and excepted matters. Reserved matters are a list of policy area—such as civil aviation, units of measurement, and human genetics, which Parliament may devolve to Northern Ireland Assembly at some time in future (Aughey, 2005). Excepted matters—such as international relations, UK taxation and elections—are never expected to be considered for devolution. On all other matters, the Northern Ireland Executive together with the 108-member Northern Ireland Assembly may legislate and govern for Northern Ireland (Aughey, 2005). In addition, devolution in Northern Ireland is dependent upon

participation by members of the Northern Ireland executive in the North/South Ministerial Council, which co-ordinates areas of co-operation—such as agriculture, education and health—between Northern Ireland and the Republic of Ireland. Elections to the Northern Ireland Assembly are by single transferable vote with six representatives elected from 18 parliamentary constituencies. Eighteen representatives to the lower house of the UK parliament are elected from the same constituencies using the first-past-the-post system. However, not all of these take their seats. In addition, the upper house of the UK parliament, the House of Lords, currently has some 25 appointed members from Northern Ireland. The Northern Ireland Office represents the UK government in Northern Ireland on reserved matters and represents Northern Ireland within the UK Government. The Northern Ireland Office is led by the Secretary of State for Northern Ireland, who sits in the UK Cabinet (Aughey, 2005).

Because of the devolution, I gave to Northern Ireland a score of 3 from 1999 on.

Scotland.—Scotland has partial self-government within the United Kingdom as well as representation in the UK Parliament. From 1999, executive and legislative powers have been devolved to the Scottish Government and the Scottish Parliament in Edinburgh, respectively (Hearn, 2002). The UK Parliament retains power over a set list of areas explicitly specified in the Scotland Act 1998 as reserved matters, including, for example, levels of UK taxes, social security, defence, international relations and broadcasting (Hearn, 2002). The Scottish Parliament has legislative authority for all other areas accounting for approximately around 70% of total identifiable public sector expenditures, as well as limited power to vary income tax (Hearn, 2002). The Scottish Parliament can give legislative consent over devolved matters back to Westminster by passing a Legislative Consent Motion (Hearn, 2002).

To take into account the devolution, I assigned to Scotland a score of 3 from 1999 on.

Wales.—Wales is a country that is part of the United Kingdom. Constitutionally, the UK is a de jure, unitary state, its parliament and government in Westminster. 40 out of the 650 representatives of the House of Commons come from Welsh constituencies. A Secretary of State for Wales sits in the UK cabinet and is responsible for representing matters pertaining to Wales. The referendum held in 1997 chose to establish a form of self-government. The consequent process of devolution began with the Government of Wales Act 1998, which created the National Assembly for Wales. Powers of the Secretary of State for

Wales were transferred to the devolved government on 1 July 1999, granting the Assembly responsibility to decide how the Westminster government’s budget for devolved areas is spent and administered (Davies et al., 2008). The 1998 Act was amended by the Government of Wales Act 2006 which enhanced the Assembly’s powers, giving it legislative powers akin to the Scottish Parliament and Northern Ireland Assembly (Davies et al., 2008). The Assembly consists of 60 members, elected for four-year terms under an additional member system. The Assembly must elect a First Minister, who selects ministers to form the Welsh Government.

The twenty areas of responsibility devolved to the Welsh Government, known as “subjects”, include agriculture, economic development, education, health, housing, local government, social services, tourism, transport and the Welsh language (Davies et al., 2008). A referendum on extending the law-making powers of the National Assembly was accordingly held on 3 March 2011. It asked the question: “Do you want the Assembly now to be able to make laws on all matters in the 20 subject areas it has powers for?” The result of the vote was that 63.49% voted “yes”, and 36.51% voted “no”. Hence, the Assembly is now able to make laws, known as Acts of the Assembly, on all matters in the subject areas, without needing the UK Parliament’s agreement on the final implementation (Davies et al., 2008).

To take into account the devolution, I assigned to Wales a score of 3 from 1999 on.

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Tables

Table I: Summary of Variables

| | Variable | Definition and Sources | Statistics |
|----------------------------|-----------------------------|---|------------------------|
| Economic outcomes: | <i>Income-1995:</i> | Natural logarithm of the GDP per capita in 1995 US dollars at the 1995 purchasing power parity exchange rates. Source: http://gecon.yale.edu/ | 9.455 (0.700) |
| | <i>Control:</i> | Unconditional average of the responses to the question “some people feel they have completely free choice and control over their lives, while other people feel that what we do has no real effect on what happens to them. Please use this scale (from 1 to 10) [...] to indicate how much freedom of choice and control in life you have over the way your life turns out.” Source: European Value Study, GESIS (2008). | 6.306 (0.709) |
| Institutions: | <i>Obedience:</i> | Share of answers mentioning “obedience” as an important quality that children should be encouraged to learn. Source: European Value Study, GESIS (2008). | 0.284 (0.133) |
| | <i>Hard-Work:</i> | Share of answers mentioning “hard work” as an important quality that children should be encouraged to learn. Source: European Value Study, GESIS (2008). | 0.438 (0.240) |
| | <i>Thrift:</i> | Share of answers mentioning “thrift” as an important quality that children should be encouraged to learn. Source: European Value Study, GESIS (2008). | 0.390 (0.117) |
| | <i>Culture-T:</i> | See text. Source: European Value Study, GESIS (2008). | 0.096 (0.347) |
| | <i>Culture-A:</i> | See text. Source: European Value Study, GESIS (2008). | 0.191 (0.443) |
| | <i>Political-Autonomy:</i> | See text. Source: Author’s codification. | 0.304 (0.535) |
| | <i>Democracy-2000-2010:</i> | <i>Democracy</i> averaged between 2000 and 2010. Sources: Author’s codification and Polity IV data set, available at http://www.systemicpeace.org | 7.620 (1.280) |
| | <i>Democracy-F:</i> | First principal component extracted from the Polity IV constraints on the executive authority score and the regional political autonomy indicator each averaged between 1950 and 2010. Sources: Author’s codification and Polity IV data set, available at http://www.systemicpeace.org | 5.982 (1.426) |
| | <i>Law-Enforcement:</i> | First principal component extracted from 2010 measures of honesty, impartiality, and quality of law enforcement and averaged over the NUTS 2 regions to which the grid cell belongs. Source: Charron et al. (2013). | 0.013 (0.992) |
| | <i>Honesty:</i> | Extent to which public education, health care, and law enforcement are shielded from corruption averaged over the NUTS 2 regions to which the grid cell belongs. Source: Charron et al. (2013). | 0.193 (0.742) |
| Quality of governance: | <i>Governance:</i> | Composite quality of governance index score rescaled in order to range between 0 and 100 and averaged over the NUTS 2 regions to which the grid cell belongs. Source: Charron et al. (2013). | 65.650 (18.647) |
| Other confounding factors: | <i>Temperature:</i> | Temperature in degree Celsius averaged over the 1961-1990 period. Source: http://gecon.yale.edu/ | 10.169 (3.423) |
| | <i>Precipitation:</i> | Precipitation in mm averaged over the 1961-1990 period. Source: http://gecon.yale.edu/ | 954.413 (313.521) |
| | <i>Land-Quality:</i> | Land quality for agriculture, defined as the probability that the grid cell may be cultivated, averaged over the grids used in the Atlas of the Biosphere and to which the grid cell belongs. Source: http://www.sage.wisc.edu/ | 0.620 (0.200) |
| | <i>Land-Quality-SD:</i> | Standard deviation of the land quality for agriculture, defined as the probability that the grid cell may be cultivated, averaged over the grids used in the Atlas of the Biosphere and to which the grid cell belongs. Source: http://www.sage.wisc.edu/ | 0.199 (0.085) |
| | <i>Area-of-Grid:</i> | Land area of the grid cell in square km. Source: http://gecon.yale.edu/ | 4769.189 (3237.387) |

Note: 1. The last column reports the mean and, in parentheses, the standard deviation of each variable. Both are computed building on the samples used in tables II, III, V, and VI.

Table II: Alternative Persistent Cultural Norms

| | (1) | (2) | (3) | (4) |
|------------------------|------------------------------------|----------------------|-----------------------|-----------------------|
| | Panel A. The dependent variable is | | | |
| | <i>Control</i> | <i>Control</i> | <i>Obedience</i> | <i>Obedience</i> |
| <i>Culture-M</i> | 0.008 (0.015) | | - 0.007 (0.003)** | |
| <i>Climate-M</i> | | 0.232 (0.199) | | - 0.162 (0.038)*** |
| <i>Ruggedness</i> | 0.625 (0.156)*** | 0.565 (0.165)*** | - 0.075 (0.030)** | - 0.033 (0.032) |
| <i>Coast</i> | - 0.049 (0.045) | - 0.043 (0.045) | - 0.016 (0.009)* | - 0.020 (0.009)** |
| Estimation | OLS | | | |
| Within R ² | 0.15 | 0.15 | 0.09 | 0.11 |
| Number of observations | 578 | 578 | 578 | 578 |
| | (1) | (2) | (3) | (4) |
| | Panel B. The dependent variable is | | | |
| | <i>Hard-Work</i> | <i>Hard-Work</i> | <i>Thrift</i> | <i>Thrift</i> |
| <i>Culture-M</i> | - 0.014 (0.003)*** | | - 0.014 (0.003)*** | |
| <i>Climate-M</i> | | - 0.089 (0.039)** | | - 0.217 (0.040)*** |
| <i>Ruggedness</i> | - 0.016 (0.030) | 0.006 (0.032) | - 0.194 (0.031)*** | - 0.138 (0.033)*** |
| <i>Coast</i> | 0.024 (0.009)*** | 0.022 (0.009)** | 0.030 (0.009)*** | 0.024 (0.009)*** |
| Estimation | OLS | | | |
| Within R ² | 0.08 | 0.05 | 0.18 | 0.19 |
| Number of observations | 578 | 578 | 578 | 578 |

Notes: 1. Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
2. All specifications also consider *Latitude*, *Longitude*, and country fixed effects.

Table III: Cross-Validating the Proxy for the Inclusiveness of Political Institutions

| | (1) | (2) | (3) |
|---------------------------|----------------------------|---------------------|---------------------|
| | The dependent variable is: | | |
| | <i>Law-Enforcement</i> | <i>Honesty</i> | <i>Governance</i> |
| <i>Political-Autonomy</i> | 0.140 (0.047)*** | 0.154 (0.018)*** | 2.777 (0.838)*** |
| Estimation | OLS | | |
| Within R ² | 0.01 | 0.02 | 0.01 |
| Number of observations | 547 | 547 | 563 |

Notes: 1. Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
2. All specifications also consider country fixed effects.

Table IV: Institutions and Outcomes — Analyzing the Overall and Regional Variation

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|---|---|---------------------|--------------------|--------------------|--------------------|--------------------|------------------|------------------|
| | The dependent variable is <i>Income</i> | | | | | | | |
| <i>Culture</i> | 1.402 (0.252)*** | 0.927 (0.123)*** | - 0.199 (0.589) | - 0.199 (0.689) | 0.127 (0.305) | 0.190 (0.414) | 0.764 (1.572) | 0.764 (1.425) |
| <i>Democracy</i> | 0.187 (0.111)* | 0.085 (0.056) | - 0.122 (0.124) | - 0.122 (0.150) | - 0.188 (0.255) | - 0.184 (0.296) | 0.710 (2.302) | 0.710 (2.086) |
| Alternative channels | YES | YES | YES | YES | YES | YES | YES | YES |
| Modulators of institutions | NO | YES | YES | YES | NO | YES | YES | YES |
| Intermediate outcomes | NO | NO | YES | YES | NO | YES | YES | YES |
| Fixed effects | NO | NO | NO | NO | YES | YES | YES | YES |
| Sanderson-Windmeijer test p-value in the first-stage for <i>Culture</i> | 0.00 | 0.00 | 0.40 | | 0.14 | 0.27 | 0.96 | |
| Sanderson-Windmeijer test p-value in the first-stage for <i>Democracy</i> | 0.00 | 0.00 | 0.36 | | 0.51 | 0.62 | 0.96 | |
| Estimation | 2SLS | 2SLS | 2SLS | 3SLS | 2SLS | 2SLS | 2SLS | 3SLS |
| P-value of underidentification test | 0.00 | 0.00 | 0.41 | | 0.46 | 0.56 | 0.95 | |
| P-value of Sargan statistic | 0.10 | 0.52 | 0.97 | 0.97 | 0.54 | 0.78 | 0.77 | 0.99 |
| Number of observations | 573 | 512 | 500 | 500 | 89 | 86 | 84 | 84 |

- Notes: 1. Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
2. All specifications also consider *Latitude* and *Longitude*. While the alternative channels are *Primary-Sector*, *Atlantic-Trade*, *LPD-M*, *Climate*, *Distance-to-Coast*, and *Traveling-Distance*, the intermediate outcomes are *Real-Capital*, *Human-Capital*, and *Catholicism*. Finally, the modulators of institutions are *Neolithic*, *Migratory-Distance*, *Potato*, and *Black-Death*. The control variables used in the second-stage are also included in the first-stage.
3. In columns (1) to (3) and (5) to (7) (columns (4) and (8)), the endogenous variables are *Culture* and *Democracy* (*Income*, *Culture*, and *Democracy*) and the excluded instruments are *Climate-M*, *Ruggedness*, and *Coast*.
4. The null hypothesis of the Sanderson-Windmeijer F test is that the endogenous variable is unidentified, and that of the Anderson underidentification (Sargan) test is that the excluded instruments are uncorrelated with the endogenous variables (exogenous).

Table V: Institutions and an Alternative Measure of Outcomes

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|--|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|
| Panel A. The dependent variable is <i>Income-1995</i> | | | | | | | | |
| <i>Culture</i> | 0.127 (0.030)*** | 0.756 (0.167)*** | 0.758 (0.165)*** | 0.745 (0.153)*** | 0.839 (0.179)*** | 0.507 (0.460) | 0.888 (0.216)*** | 0.878 (0.318)*** |
| <i>Democracy</i> | 0.085 (0.013)*** | 0.084 (0.122) | 0.083 (0.121) | 0.089 (0.118) | 0.173 (0.140) | 0.451 (0.558) | 0.025 (0.118) | 0.016 (0.186) |
| Sanderson-Windmeijer test p-value in the first-stage for <i>Culture</i> | | 0.00 | 0.00 | 0.00 | 0.00 | 0.66 | 0.00 | 0.04 |
| Sanderson-Windmeijer test p-value in the first-stage for <i>Democracy</i> | | 0.00 | 0.00 | 0.00 | 0.01 | 0.71 | 0.00 | 0.09 |
| Estimation | OLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS |
| Within R ² | 0.15 | | | | | | | |
| P-value of underidentification test | | 0.00 | 0.00 | 0.00 | 0.01 | 0.71 | 0.00 | 0.10 |
| P-value of Sargan statistic | | 0.32 | 0.33 | 0.31 | 0.34 | 0.30 | 0.73 | 0.42 |
| Number of observations | 578 | 578 | 573 | 578 | 578 | 578 | 578 | 578 |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Panel B. The dependent variable is <i>Income-1995</i> | | | | | | | | |
| <i>Culture</i> | 0.636 (0.290)** | 0.747 (0.182)*** | 0.756 (0.169)*** | 0.937 (0.265)*** | 0.880 (0.196)*** | 0.746 (0.165)*** | 0.066 (0.036)* | 0.737 (0.267)*** |
| <i>Democracy</i> | 0.119 (0.134) | 0.093 (0.120) | 0.074 (0.129) | - 0.184 (0.154) | 0.210 (0.181) | 0.116 (0.136) | 0.044 (0.013)*** | 0.122 (0.161) |
| Sanderson-Windmeijer test p-value in the first-stage for <i>Culture</i> | 0.02 | 0.00 | 0.00 | 0.00 | 0.00 | 0.00 | | 0.00 |
| Sanderson-Windmeijer test p-value in the first-stage for <i>Democracy</i> | 0.02 | 0.00 | 0.00 | 0.02 | 0.04 | 0.01 | | 0.08 |
| Estimation | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | 2SLS | OLS | 2SLS |
| Within R ² | | | | | | | 0.48 | |
| P-value of underidentification test | 0.03 | 0.00 | 0.00 | 0.02 | 0.04 | 0.01 | | 0.07 |
| P-value of Sargan statistic | 0.19 | 0.26 | 0.32 | 0.55 | 0.33 | 0.22 | | 0.31 |
| Number of observations | 563 | 578 | 578 | 518 | 577 | 578 | 500 | 500 |

- Notes:
1. Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
 2. All specifications also consider *Latitude*, *Longitude*, and country fixed effects. The specifications in columns (3) to (8) of panel A also include *Primary-Sector*, *Atlantic-Trade*, *LPD-M*, *Climate*, *Distance-to-Coast*, and *Traveling-Distance*, respectively, whereas those in columns (1) to (6) of panel B also incorporate *Real-Capital*, *Human-Capital*, *Catholicism*, *Neolithic* and *Migratory-Distance*, *Potato*, and *Black-Death*, respectively. The extra controls considered in the specifications in columns (7) to (9) of panel B are *Primary-Sector*, *Atlantic-Trade*, *LPD-M*, *Climate*, *Distance-to-Coast*, *Traveling-Distance*, *Real-Capital*, *Human-Capital*, *Catholicism*, *Neolithic*, *Migratory-Distance*, *Potato*, and *Black-Death*. The control variables used in the second-stage are also included in the first-stage.
 3. In columns (2) to (9) of panel A and columns (1) to (6) and (8) of panel B (column (9) of panel B), the endogenous variables are *Culture* and *Democracy* (*Income*, *Culture*, and *Democracy*) and the excluded instruments are *Climate-M*, *Ruggedness*, and *Coast*.
 4. The null hypothesis of the Sanderson-Windmeijer F test is that the endogenous variable is unidentified, and that of the Anderson underidentification (Sargan) test is that the excluded instruments are uncorrelated with the endogenous variables (exogenous).

Table VI: Alternative Measures of Institutions, Controls, and Instruments

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
|-------------------------------------|--|--------------------|---------------------|---------------------|---------------------|---------------------|---------------------|---------------------|----------------------|---------------------|
| | The dependent variable is <i>Income</i> | | | | | | | | | |
| <i>Trust</i> | 1.160 (0.396)*** | | | | | | | | | |
| <i>Respect</i> | | 1.944 (0.854)** | | | | | | | | |
| <i>Culture-T</i> | | | 0.566 (0.177)*** | | | | | | | |
| <i>Culture-A</i> | | | | 0.432 (0.123)*** | | | | | | |
| <i>Culture</i> | | | | | 0.526 (0.194)*** | 0.568 (0.192)*** | 0.947 (0.298)*** | 0.732 (0.174)*** | 0.728 (0.197)*** | 3.262 (16.553) |
| <i>Democracy</i> | 0.001 (0.108) | 0.098 (0.152) | 0.049 (0.108) | - 0.003 (0.094) | | | 0.138 (0.185) | 0.232 (0.144) | 0.225 (0.177) | - 5.404 (26.586) |
| <i>Democracy-2000-2010</i> | | | | | 0.044 (0.151) | | | | | |
| <i>Democracy-F</i> | | | | | | 0.044 (0.116) | | | | |
| <i>Temperature</i> | | | | | | | 0.018 (0.015) | | | |
| <i>Precipitation</i> | | | | | | | 0.0001 (0.0001) | | | |
| <i>Land-Quality</i> | | | | | | | | 0.322 (0.132)** | | |
| <i>Land-Quality-SD</i> | | | | | | | | 0.506 (0.189)*** | | |
| <i>Area-of-Grid</i> | | | | | | | | | 4.79E-6 (6.99E-6) | |
| | P-value of the Sanderson-Windmeijer test in the first-stage for the proxy for a culture of cooperation | | | | | | | | | |
| | 0.00 | 0.04 | 0.00 | 0.00 | 0.02 | 0.00 | 0.00 | 0.00 | 0.01 | 0.84 |
| | P-value of the Sanderson-Windmeijer test in the first-stage for the proxy for inclusive political institutions | | | | | | | | | |
| | 0.07 | 0.12 | 0.08 | 0.07 | 0.17 | 0.08 | 0.07 | 0.02 | 0.10 | 0.84 |
| Estimation | 2SLS | | | | | | | | | |
| P-value of underidentification test | 0.06 | 0.12 | 0.07 | 0.07 | 0.17 | 0.07 | 0.07 | 0.03 | 0.11 | 0.84 |
| P-value of Sargan statistic | 0.32 | 0.39 | 0.73 | 0.64 | 0.68 | 0.75 | 0.35 | 0.45 | 0.46 | |
| Number of observations | 500 | 500 | 500 | 500 | 500 | 500 | 578 | 578 | 578 | 500 |

- Notes:
- Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
 - All specifications also consider *Latitude*, *Longitude*, and country fixed effects. The extra controls included in the specifications in columns (1) to (6) and (10) are *Primary-Sector*, *Atlantic-Trade*, *LPD-M*, *Climate*, *Distance-to-Coast*, *Traveling-Distance*, *Real-Capital*, *Human-Capital*, *Catholicism*, *Neolithic*, *Migratory-Distance*, *Potato*, and *Black-Death*. The control variables used in the second-stage are also included in the first-stage.
 - The endogenous variables are *Trust* and *Democracy* in column (1), *Respect* and *Democracy* in column (2), *Culture-T* and *Democracy* in column (3), *Culture-A* and *Democracy* in column (4), *Culture* and *Democracy-2000-2010* in column (5), *Culture* and *Democracy-F* in column (6), and *Culture* and *Democracy* in columns (7) to (10). The excluded instruments are *Climate-M*, *Ruggedness*, and *Coast* in columns (1) to (9) and *Culture-M* and *Democracy-M* otherwise.
 - The null hypothesis of the Sanderson-Windmeijer F test is that the endogenous variable is unidentified, and that of the Anderson underidentification (Sargan) test is that the excluded instruments are uncorrelated with the endogenous variables (exogenous).

Table VII: Institutions and Outcomes — Excluding Outliers

| | (1) | (2) | (3) | (4) |
|-------------------------------------|----------------------------|---------------------|--------------------|---------------------|
| | The dependent variable is: | | | |
| | <i>Income</i> | <i>Culture</i> | <i>Democracy</i> | <i>Income</i> |
| <i>Culture</i> | 0.312 (0.026)*** | | | 1.049 (0.235)*** |
| <i>Democracy</i> | 0.047 (0.009)*** | | | - 0.230 (0.182) |
| <i>Climate-M</i> | | 0.343 (0.105)*** | - 0.010 (0.301) | |
| <i>Ruggedness</i> | | 0.279 (0.088)*** | 0.561 (0.253)** | |
| <i>Coast</i> | | - 0.042 (0.024)* | 0.100 (0.070) | |
| Estimation | OLS | OLS | OLS | 2SLS |
| Within R ² | 0.31 | | | |
| Sanderson-Windmeijer test p-value | | 0.00 | 0.06 | |
| P-value of underidentification test | | | | 0.06 |
| P-value of Sargan statistic | | | | 0.97 |
| Number of observations | 546 | 546 | 546 | 546 |

- Notes:
- Standard errors in parentheses. *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
 - All specifications also consider *Latitude*, *Longitude*, and country fixed effects. The control variables used in the second-stage are also included in the first-stage.
 - The endogenous variables are *Culture* and *Democracy*, and the excluded instruments are *Climate-M*, *Ruggedness*, and *Coast*.
 - The null hypothesis of the Sanderson-Windmeijer F test is that the endogenous variable is unidentified, and that of the Anderson underidentification (Sargan) test is that the excluded instruments are uncorrelated with the endogenous variables (exogenous).

Table VIII: Allowing for Clustering by Country and Dealing With Spatial Correlation

| | (1) | (2) | (3) | (4) | (5) | (6) |
|-------------------------------------|-----------------------|---------------------|----------------------------|----------------------|--------------------|---------------------|
| | <i>Culture</i> | <i>Democracy</i> | The dependent variable is: | | <i>Democracy</i> | <i>Income</i> |
| <i>Culture</i> | | | 0.779 (0.378)** | | | 0.788 (0.243)*** |
| <i>Democracy</i> | | | 0.176 (0.188) | | | 0.156 (0.189) |
| <i>Climate-M</i> | 0.398 (0.172)** | - 0.025 (0.510) | | 0.398 (0.130)*** | - 0.025 (0.322) | |
| <i>Ruggedness</i> | 0.155 (0.163) | 0.688 (0.241)*** | | 0.155 (0.136) | 0.688 (0.312)** | |
| <i>Coast</i> | - 0.101 (0.033)*** | 0.112 (0.082) | | - 0.101 (0.049)** | 0.112 (0.100) | |
| Sanderson-Windmeijer test p-value | 0.01 | 0.02 | | | | |
| Estimation | OLS | OLS | 2SLS | OLS | OLS | 2SLS |
| Sanderson-Windmeijer test p-value | 0.01 | 0.02 | | | | |
| P-value of underidentification test | | | 0.14 | | | |
| P-value of Sargan statistic | | | 0.70 | | | 0.32 |
| Number of observations | 578 | 578 | 578 | 578 | 578 | 578 |

- Notes:
1. Robust standard errors allowing for clustering by country (Conley's (1999) standard errors) in the parentheses of columns (1) to (3) (columns (4) to (6)). *** denotes significant at the 1% confidence level; **, 5%; *, 10%.
 2. All specifications also consider *Latitude*, *Longitude*, and country fixed effects. The control variables used in the second-stage are also included in the first-stage.
 3. The endogenous variables are *Culture* and *Democracy*, and the excluded instruments are *Climate-M*, *Ruggedness*, and *Coast*.
 4. The null hypothesis of the Sanderson-Windmeijer F test is that the endogenous variable is unidentified, and that of the Anderson underidentification (Sargan) test is that the excluded instruments are uncorrelated with the endogenous variables (exogenous).